Impacts of hot-dry compound extremes on US soybean yields

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Abstract. The US agriculture system supplies more than one-third of globally-traded soybean and with 90% of US soybean produced under rainfed agriculture, soybean trade is particularly sensitive to weather and climate variability. Average growing season climate conditions can explain about one-third of US soybean yield variability. Additionally, crops can be sensitive to specific short-term weather extremes, occurring in isolation or compounding at key moments throughout crop development.

- 5 Here, we identify the dominant within-season climate drivers that can explain soybean yield variability in the US, and explore synergistic effects between drivers that can lead to severe impacts. The study combines weather data from reanalysis and -, satellite-informed -based evapotranspiration and root-zone soil moisture fields with sub-national crop yields using statistical methods that account for interaction effects. Our models can explain on average about half-two thirds of the year-to-year yield variability (6070% on all years and 4060% on out-of-sample predictions). The largest negative influence on soybean yields is
- 10 driven by high temperature and low soil moisture during the summer crop reproductive period. Moreover, due to synergistic effects, heat is considerably more damaging to soybean crops during dry conditions, and less so during wet conditions. Compound and interacting hot and dry <u>August-summer</u> conditions (defined by the 95th and 5th percentiles of temperature and soil moisture, respectively) reduce yields by <u>1.252</u> standard deviation. This sensitivity is, respectively, <u>46</u> and 3 times larger than the sensitivity to hot or dry conditions alone. Other <u>important-relevant</u> drivers of negative yield responses are lower
- 15 evapotranspiration temperatures early and late in the season, excessive precipitation in early season and dry conditions in late season and lower minimum temperature late in the season, both likely reflecting an increased risk of frost... The sensitivity to the identified drivers varies across the spatial domain with higher latitudes, and thus colder regions, positively affected by high temperature during the summer period. On the other hand,- and-warmer south-eastern regions are positively affected by low temperature late season, being less sensitive to hot dry August months... Historic trends in identified drivers indicates that US
- 20 soybean has generally benefited from recent shifts in weather except for increasing rainfall in the early season. Overall warming conditions have reduced the risk of frost in early and late-season and potentially allowed for earlier sowing dates. More importantly, summers have been getting cooler and wetter over eastern US. Still, despite these positive changes, we show that the frequency of compound hot-dry August monthsummer events has remained unchanged over 1946-2016. Moreover, inIn the longer term, climate models project substantially warmer summers for the continental US but uncertainty remains whether
- 25 this will be accompanied by drier conditions. This highlightsing a critical element to explore in future studies focused on US agricultural production risk under climate change. -which likely creates risks for soybean production.

1 Introduction

Soybean is one of the most in-demand crops worldwide, with the largest increases in production-area over the last two decades

- 30 when compared to all other major staple crops (Hartman et al., 2011).-<u>A considerably large portion of this production is</u> <u>dedicated to animal feed accommodating the current global increase in demand for animal products</u> (Cassidy et al., 2013).<u>A</u> recent estimate based on FAOSTAT data in 2013 reports that soybean ranks second in terms of globally-produced kilocalories (~20% of the total kcal traded on the global food market) and first among staple crops in terms of globally-aggregated trade monetary value (Torreggiani et al., 2018). The US agriculture system alone supplies more than one-third of globally-traded
- 35 soybean, of which 90% is produced under rainfed agriculture (Jin et al., 2017). The recent surge in global soybean demands is expected to increase further in the future due to increasing global population and associated shifts in dietary preferences (Fehlenberg et al., 2017). At the same time, climate change is expected to increase annual mean and extreme temperature levels over the US (Dirmeyer et al., 2013; Winter et al., 2015; Wuebbles et al., 2014a). To support adaptation measures that reduce the potential impacts of these future challenges, we need a quantitative understanding of crop sensitivity to climate and weather

40 variables.

Climate variability can strongly impact crop yields. The effects of growing season temperature and precipitation conditions can explain about one-third of US soybean year-to-year yield variability (Leng et al., 2016; Lobell et al., 2011; Ray et al., 2015; Vogel et al., 2019). In particular, heat and drought conditions are among the most limiting environmental factors

- 45 affecting crops (Lesk et al., 2016). These are increasingly detrimental when coinciding with vulnerable stages of the crop growth cycle (Troy et al., 2015). Such conditions can occur separately or in combination, in the latter case, leading often to more severe impacts (Leonard et al., 2014). For instance, it is reported that US economic agricultural losses between 1980 and 2012 are four times larger during hot and dry conditions compared to drought events alone (Suzuki et al., 2014). Moreover, the response to multiple climatic stressors is complex and can be subject to interaction effects where climatic drivers create
- 50 more damage in combination than the sum of each in isolation (Ben-Ari et al., 2018; Haqiqi et al., 2021; Matiu et al., 2017; Rigden et al., 2020). Interestingly, multiple climatic stressors can also result in positive interactions with beneficial effects on crop yields (Carter et al., 2016; Suzuki et al., 2014). Such features, positive or negative, are likely to have important implications on future impacts and adaptation strategies to climate change. Nevertheless, these have received little attention in current assessments so far (Matiu et al., 2017; Zscheischler et al., 2017).
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A compound event framework has lately been proposed to underline the need for impact-centric approaches that identify multiple climatic drivers contributing to socio-economic risk (Leonard et al., 2014; Zscheischler et al., 2018, 2020). The types of damaging combination of drivers on local agricultural production are various, with a specific terminology recently proposed

in Zscheischler et al. (2020). These can be temporally compounding, as in the case of the 2016 wheat production in France

- 60 where high temperatures during winter followed by heavy precipitation during spring lead to unprecedented yield losses (Ben-Ari et al., 2018). These can be preconditioned where for instance, pre-sowing soil moisture water storage content interacts with within-season precipitation to affect rainfed maize yield in the US (Carter et al., 2018a) or multivariate/co-occurring such as in the case of hot-dry conditions in the growing season affecting crop yields (Feng and Hao, 2019; Matiu et al., 2017). One way to identify such drivers is through the use of statistical methods that empirically associate drivers to impacts (Vogel et al.,
- 65 2021). Easily interpretable linear regressions in that context can be useful tools, in particular when fitted with alternative methods that allow for the consideration of a large number of potential predictors (i.e. subset selection, shrinkage or dimension reduction approaches) (Ben-Ari et al., 2018; Carter et al., 2018a; Laudien et al., 2020; Vogel et al., 2021).

Here we analyze soybean yields and climate time series for the U.S. at the county scale from 1982 to 2016 using regression models that are fitted with a reduced set of variables selected via a subset selection approach. The aim is to identify (1) the combination of climatic conditions affecting soybean yields at different stages of the growing season, and (2) potential interaction effects between drivers modulating the final impact on yield. <u>Although other studies have looked at potential</u> interactions between climate drivers, identifying key month and variables throughout the growing season, these studies did not look at such features jointly as done here. Furthermore, we study (3) trends in the identified dominant climate drivers from

75 1946 to 2016 to assess how historic trends likely affected current-soybean production risk. Finally, we explore (4) how temperature and moisture couplings differ within the growing season between hot-dry summers and normal summers. We discuss how that potentially affects the occurrence of compound hot-dry extremes and associated crop impacts. Finally, we investigate trends in the identified dominant climate drivers from 1946 to 2016 to assess how historic trends have affected current soybean production risk.

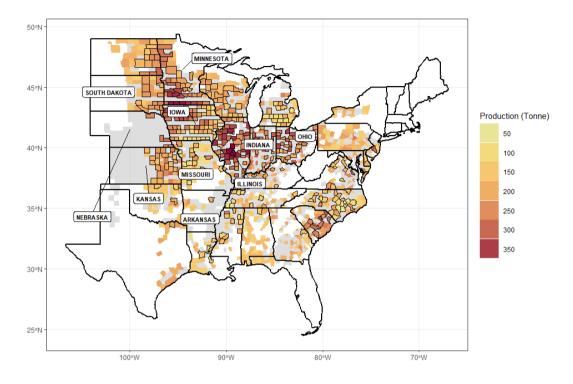
80 2 Data and Methods

2.1 Soybean yields, and climate and hydrological data for the U.S.

Soybean yields are analysed at the county scale for the period 1982-2016, based on census data obtained from the US Department of Agriculture (USDA) National Agriculture Statistics Survey (NASS) Quick Stats database (www.nass.usda.gov/Quick Stats). Counties are selected on (i) having no missing data for the full 35 years analysed, (ii) have common planting dates (i.e. April-May) and (iii) a production area share of at least 90% rainfed agriculture. Consequently, a total of 389 counties are retained for the regression analysis (Fig. 1). These together account for at least 50% of US total rainfed soy production, where production per county is calculated as the average production over 1982-2016. In the study region, planting dates are aligned to provide comparable crop growth stages between counties. This facilitates the interpretation of climate sensitivities associated to timing within the growing season.-Information on the soybean growing season and rainfed

90 vs irrigated agricultural land cover is obtained from the monthly irrigated and rainfed crop areas database around the year 2000

(MIRCA2000), a global gridded dataset at 0.5° resolution (Portmann et al., 2010). The percent rainfed area is calculated by dividing the rainfed area in each grid cell by the total harvested area for each cell (Schauberger et al., 2017a). A linear trend is removed from yield values at the county scale to eliminate long-term effects largely due to technological improvements over the study period (Fig. S1) -(Li et al., 2019; Zipper et al., 2016).



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Figure 1. Average total production in tonnes over the period of study (1982-2016). Counties with 35 years of data are highlighted with a thin black perimeter. Grey regions represent filtered out counties where local agriculture is less than 90% rainfed.

Global hydrological and weather datasets are used for this analysis. This provides the possibility to conduct similar assessments, in other parts of the world, whenever impact data is available. Nevertheless, other studies can benefit from
 leveraging local climate and hydrological data when available for better representativeness. Root zone soil moisture (SMroot) variable (m³/m³) is obtained from the modelled GLEAM v3.3a dataset that incorporates an observed satellite-based soil moisture data assimilation system (Martens et al., 2017). The dataset is available at a 0.25° grid resolution and a daily time step that covers the period of study (1982-2016). Weather data, namely maximum (Tmax) and minimum (Tmin) temperature (°C) in addition to precipitation (mm) are obtained from the bias-adjusted WFDE5 reanalysis covering the same period (1982-2016) at daily time step and a 0.5° grid resolution (Cucchi et al., 2020). Daily precipitation is further processed into number of days with precipitation above 20 mm (Num pr20) to explicitly account for potential negative effects of excessive precipitation on yield (Li et al., 2019; Zhu and Troy, 2018). All variables are temporally aggregated to monthly and seasonal

windows over springearly- (April-May), summermid- -(June-July-August) and autumnlate-growing season (September-

October) periods. Additionally, variables are spatially aggregated to the county scale based on county boundary maps of the

- 110 <u>2016 US Census Bureau.</u> Climate data are obtained from the bias adjusted WFDE5 global reanalysis covering the same period (1982-2016) on a daily time step at a 0.5° grid resolution (Cucchi et al., 2020). Monthly values in each grid cell are calculated for the following variables: the monthly mean daily maximum (Tmax) and minimum temperatures (Tmin) (°C), monthly mean precipitation (mm), cumulative incident solar radiation (Wm⁻²) in addition to extreme indicators such as number of days with temperature above 30 °C (i.e. soybean critical temperature threshold) (Schlenker and Roberts, 2009), and number of days
- 115 with precipitation above 1, 20, and 30 mm to account for potential negative effects of excessive precipitation on yield (Li et al., 2019). Additional variables are created by aggregating over the spring (April May), summer (June July August) and autumn (September October) periods. Actual evapotranspiration (mm) and root zone soil moisture (m³/m³) from the satellite-based GLEAM dataset (Martens et al., 2017) are included in the analysis at the same spatio temporal scale. All input data is then averaged over the area of each county. A summary of the considered variables for the modelling analysis is presented in
- 120 Table 1. Dividing the growing season by calendar months allowed the identification of key phases throughout the season where soybean crops are most sensitive to climate variability. These can reflect both vulnerable physiological crop growth stages and important climatic thresholds. We could have used a more complex characterization of crop developmental stages based on phenological heat units (Schauberger et al., 2017b) or the consideration of sub-monthly aggregation periods for climatic time series, but these did not necessarily improve model performance in other assessments and therefore we opted here to simply
- 125 rely on monthly and seasonal estimates (Ben-Ari et al., 2016; Ortiz-Bobea et al., 2019; Sharif et al., 2017). <u>Full growing season</u> averages have been tested as potential predictors but these did not improve modelling results and have therefore been omitted from further analysis. We thus exclusively focus on within season crop climate sensitivities.

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Table 1. Climate variables calculated at seasonal and monthly time scales throughout the growing season

Variable abbreviation	Variable explanation	<u>Type</u>	<u>Unit</u>
<u>Tmin</u>	Average minimum Temperature	Temperature related	<u>°C</u>
<u>Tmax</u>	Average maximum Temperature	Temperature related	<u>°C</u>
Num pr20	Number of days with precipitation above 20 mm	Moisture related	<u>days</u>
<u>SMroot</u>	Root zone soil moisture	Moisture related	$\underline{m^3/m^3}$

	ſ	Variable abbreviation	Variable explanation	Unit
4		rsds	Shortwave radiation	W/m^2
Hea	-	Tmin	Average minimum Temperature	°€

	Tmax	Average maximum Temperature		
	Num_tx30	Number of days with temperature above 30 °C	days	
	Precip_avg	Average amount of precipitation	mm	
	Num_wet	Number of days with precipitation above 1 mm	days	
related	Num_pr20	Number of days with precipitation above 20 mm	days	
	Num_pr30	Number of days with precipitation above 30 mm	days	
4 oisture	SMroot	Root zone soil moisture	m^{3}/m^{3}	
Moi	ETact	Actual evapotranspiration	mm	

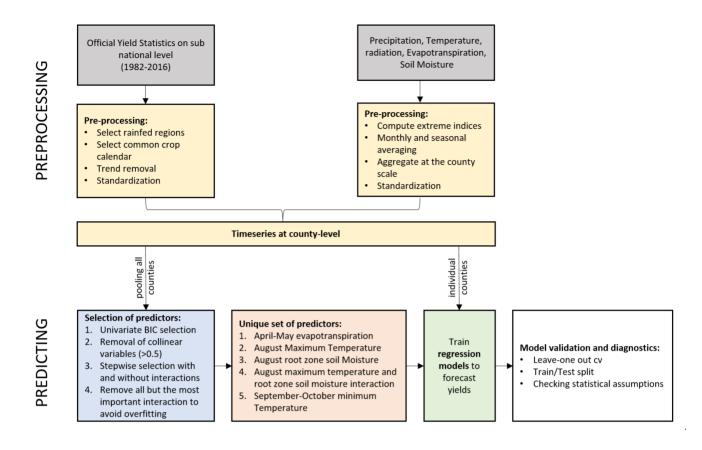
2.2 Simulating yield variability

We used regression models to estimate yield variability at the county scale. Typically, three types of statistical models are used 135 in such assessments (i.e. time-series, panel, and cross-sectional models) (Lobell and Burke, 2010). Here we opted for timeseries models as these are (i) easy to interpret, (ii) often perform well compared to the other approaches, and (iii) allow for spatially heterogeneous parameter estimation that may highlight important local and regional features (Gornott and Wechsung, 2016). To focus on robust precursors and to enhance model interpretability, we first selected one set of predictors for the full region by pooling US county yields together (see Fig. 2, box 'selection of predictors') (Troy et al., 2015). Out of all possible models constructed with a single input variable at county scale, we selected the most influential moisture- and heattemperature-140related variables per county based on the Bayesian Information Criterion (BIC) (Ben-Ari et al., 2018). We This was done separately do this for early- (April-Mayspring), mid- (June-July-Augustsummer) and late-growing season (September-Octoberautumn) periods separately considering both monthly and seasonal aggregates for each, and thus, ending ended up with a subset of six best predictors for each county. (see Table A1). To avoid multicollinearity, we pruned this list of selected 145 predictors by setting a maximum allowable Pearson correlation coefficient between any two predictors to 0.5. Thus, whenever a pair of predictors was strongly collinear (Pearson's r > 0.5), we selected the predictor that preceded the other in timing within the growing season (i.e. we excluded soil moisture in September as August soil moisture was already selected). Finally, we applied a stepwise selection procedure -to identify the best combination of these input variables, with and without interactions, picking the model with the lowest BIC value at county level (Ben-Ari et al., 2018). The stepwise approach considers all selected variables and all possible interactions (i.e. products of all possible pairs of selected predictors). The procedure is then to start 150 from a model with no predictors, sequentially adding and removing predictors until only a subset is left resulting in the most parsimonious model with the lowest prediction error on training data (See step.lm function of R, version 3.6.1). The

the final model as this was shown to reduce overfitting. The final list of selected predictors consisted of April May 155 evapotranspiration, August root zone soil moisture, August maximum temperature, September October minimum temperature, and the interaction between temperature and soil moisture in August (see Fig. 2, box 'unique set of predictors').

performance of the resulting model was -Only interactions that improved the model out of sample performance were kept in

The resulting model is fitted at the county scale and its performance is evaluated using the coefficient of determination (R²). Further robustness tests with respect to both predictor selection and model performance are detailed in the following subsection.-A summary of the modelling framework is presented in Fig. 2.



PREDICTING

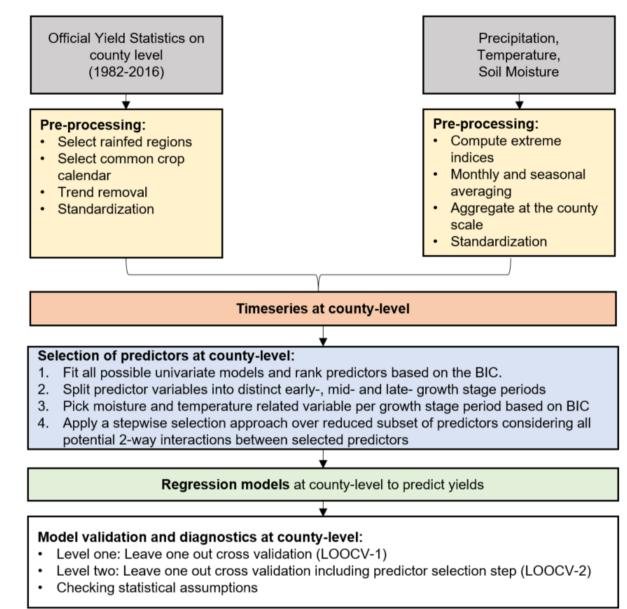


Figure 2. Overall modelling workflow applied for this study linking US yields to weather and climate variables.

2.3 Validating performance and testing modelling assumptions

To test robustness of the model performance and the selected predictors, we applied a two level leave-one-out cross-validation scheme (LOOCV) (Laudien et al., 2020). We-Level one (LOOCV-1) consisted of trainingtrained _county-scale models on reduced datasets. These are constructed by where we iteratively removinged the to-be-forecasted year and predictinged the one out of sample value using a set of predictors per county selected using the complete dataset. Level two (LOOCV-2) is similar but repeats the predictor selection step for every iteration. This way, we completely eliminate information shared between training and validation sets. Furthermore, we calculated how often selected predictors are chosen across each iteration

- 170 in the cross-validation procedure of LOOCV-2. Both elements, respectively, provide a more robust model performance estimate and predictor selection step In addition, we applied an alternative cross validation method using a train test split approach where we trained the model over the first 18 years and tested its performance over the remaining 17 years of data. The adequacy of applying linear models at the county scale for assessing the relationship between yield anomalies and selected predictors was successfully assessed using five statistical tests (Gornott and Wechsung, 2016; Schauberger et al., 2017b). The
- 175 regression equation specification error test (RESET) assessed whether taking powers of the predictor variables would improve the model fit. The Breusch-Pagan test examined heteroscedasticity issues with the data. The Breusch–Godfrey test was used to assess autocorrelation and the Shapiro–Wilk test to examine normality of residuals. Multicollinearity was checked using the variance inflation factor calculated for each independent variable while setting acceptable levels to strictly below 3.

2.4 Changes in key climatic conditions from 1946 to 2016

- 180 Historic trends of the dominant climatic drivers were assessed for the period 1946 to 2016 using linear regressions (0.05-05 significance level). Furthermore, we assessed changes in concurrent hot-dry Augustsummer conditions as these were shown to be particularly relevant for soybean production. The selected input datasets used in the crop-modelling analysis do not cover years preceding 1981. To overcome this limitation, we used precipitation, number of wet days and temperature minimum and maximum variables from the CRU V4 global dataset (Harris et al., 2020) covering the period 1901-2019 at a spatial resolution
- 185 of 0.5°. <u>Number of wet days</u> <u>-Minimum temperature</u> in the early season was used as a proxy for early season <u>number of days</u> with precipitation above 20 mm. actual evapotranspiration as the latter tends to be mainly energy limited during spring, especially in climatologically wet regions. Moreover, minimum temperature was initially picked as most relevant temperature related variable for spring conditions, but later dropped in the stepwise selection. Mean summer precipitation over June-July-August-September was used as a proxy for August-September averaged root zone soil moisture. To check the feasibility of
- 190 these assumptions, we calculated correlation maps between GLEAM August-<u>September averaged</u> root zone soil moisture and CRU averaged summer precipitation and between <u>WFDE5 spring number of days with precipitation above 20 mm</u>-<u>GLEAM</u> spring actual evapotranspiration and CRU spring <u>number of wet days</u> minimum temperature for the period 1982 to 2016. The mean Pearson's correlation coefficient over the whole spatial domain was 0.73-66 for summer precipitation and root zone soil moisture and 0.83 for spring number of wet days and number of wet days above 20 mm <u>and 0.5 for spring actual</u>
- 195 evapotranspiration and minimum temperature (Fig. A1S2ab). The 10th25th/10th –and 75th/90th percentiles of summer precipitation and August maximum temperature are used to jointly define the compound hot-dry events at the local scale. Accordingly, we calculated the percent-change per grid cell based on the difference between the number of compound events over two distinct periods (1946-1980 relative to 1982-2016) normalized by the total amount of events over the entire analysis period. Statistical significance of this percent change is assessed using the non-parametric Wilcoxon Rank Sum test (0.05).
- 200 <u>significance level</u>). <u>MM</u>oreover, we calculated a percent (%) area time series of the total rainfed producing region under

compound <u>August summer</u> hot-dry conditions by summing the number of grid cells under such conditions for a given year and dividing by the total number of grid cells considered, similar to the approach applied in Mazdiyasni and AghaKouchak (2015). The trend in the aforementioned time_-series was assessed with the non-parametric Mann–Kendall trend test (0.05 significance level).

205 2.5 Exploring temperature and moisture couplings during summer hot-dry events

To get insight on how key elements related to moisture and temperature couplings differ during compound summer hot-dry years, we estimated the coevolution of actual evapotranspiration, root zone soil moisture and maximum temperature pairs composited into hot-dry events for the period 1982-2016. Hot-dry summer events in this case are defined as years when more than 20% of the total eonsidered spatial domain harvested area is under hot-dry conditions (using the 75th and 25th percentiles respectively). Coevolution of considered variables was estimated by calculating monthly correlations across the year for all spatially averaged variable-pair combinations. Actual evapotranspiration (AET) (mm) is retrieved from the GLEAM v3.3a dataset with the same temporal and spatial resolution of aforementioned root zone soil moisture variable. AET within the GLEAM dataset is derived from potential evapotranspiration model estimates multiplied by an evaporative stress factor based on observations of microwave vegetation optical depth (VOD) and root zone soil moisture values.

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3 Results

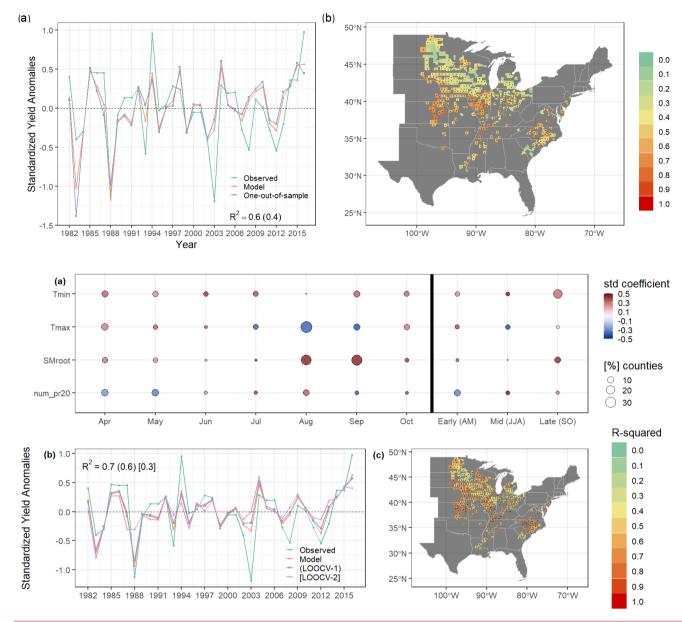
3.1 Overall model performance

Based on the selection procedure shown in Fig. 2, we identify a set of non-redundant moisture and temperature variables at different stages of the growing season that can best explain yield variability at county scale. These varied across the spatial

- 220 domain (Fig. A1, A2) with dominant patterns summarized as follows: Excessive precipitation is highlighted as the main driver of reduced soybean yields in the early season alongside low minimum and maximum temperature values. Low soil moisture and high maximum temperature values are highlighted as main drivers of reduced yields in the mid-season, particularly for the months of August and September. Finally, low soil moisture and low minimum temperature values are highlighted as main drivers of reduced yields late in the season (Fig. 3a). we identify a set of predictors for the full region consisting of (1) April-
- 225 May evapotranspiration, (2) September October minimum temperature, (3) August root zone soil moisture, (4) August maximum temperature, and (5) the interaction between those two August variables. These unique predictors represent non-redundant moisture and heat conditions at different stages of the growing season. The trained regression models are trained on theat countrty level_-with those identified predictors_and are able to explain, on average, about two-thirdshalf of the year-to-year yield variability (6070% on all years and 4060% on level_HLOOCV-1_out-of_sample-predictions). Including interaction
- 230 terms in the fitted model contributed to 10% out of the total 60% explained variability on level 1LOOCV-1 -out of sample ppredictions. –Testing the model with the more conservative level 2 cross validationLOOCV-2, repeating the predictor

selection step at every iteration, lowers model explained variability to 30% (Fig. 3a). This reduced performance is expected when comparing with results of studies that applied a similar robust leave one out cross-validation approach (Laudien et al., 2020; Lehmann et al., 2020). Still, for ~83% of the years, the LOOCV-2 model provides a correct year-to-year direction of

- 235 change as well as sign of the yield anomaly (i.e. positive or negative) (Fig. 3b). Furthermore, most frequently selected predictors and associated timing within the season across the training sets shows high consistency and good agreement with predictors selected on the full dataset (Fig. S3, S4, S5). This provides confidence with respect to the choice of predictors (Sup). The train test split approach gives quantitatively similar performance results so we limit ourselves here to presenting the leave one out cross validated results (See Fig. A2). In general, for almost all years, the model provides a correct year to year
- 240 direction of change as well as sign of the yield anomalies (i.e. positive or negative, see Fig 3a). Overall, the most importantdominant crop yield drivers are August/September root zone soil moisture and August maximum temperature, each selected over more than 25% of considered counties. Averaged standardized beta coefficients for aforementioned variables reports the highest absolute value of around 0.4 (i.e. ~0.4 standard deviation change in soybean yields per standard deviation change in the predictor when excluding the effect of interaction terms). Furthermore, interaction
- 245 <u>effects between summer moisture and temperature variables are the most frequently selected type of interaction (Fig. A3).</u> <u>together responsible for 65% of the model out of sample explained variability. Including the interaction term between those variables contributed to 12.5% out of the total 65% attributed to August heat and moisture variables. The co-occurrence of low soil moisture and hot conditions triggers the largest crop failures. Extreme hot dry conditions (i.e. simultaneously exceeding the 95th and 5th percentiles of temperature and soil moisture, respectively) leads to 6 times more crop impacts compared to</u>
- 250 extreme hot conditions alone (i.e. 95th and 50th percentiles of temperature and soil moisture, respectively) and 3 times more impacts compared to extreme dry conditions alone (i.e. 50th and 5th percentiles of temperature and soil moisture, respectively). (Fig. 3a).



255 Figure 3. (a) summary of the strength and frequency of selected predictors across the growing season. (b-c) Explained variance (R-squared) of yield anomalies due to climate variability (ab) spatially averaged and (bc) at the county scale. Stippling in (bc) shows F-tests with (p < 0.05) indicating that the model chosen is significantly better than a null model (accounting for false discovery rate due to multiple hypotheses testing).</p>

In particular, extremely low yields occurring during heat and drought events, such as the 1988 and 2012 years, are well captured by the model. <u>SS</u>patially, the model is statistically significant (p-value < 0.05) for over 81% of <u>all</u> considered counties (Fig. 3c) and 77% when we<u>after</u> adjusted <u>adjusting</u> for multiple hypotheses testing using the False Discovery Rate (FDR) method (Ventura et al., 2004). Yield variability is captured particularly well in southern counties (Fig. 3b3c), with high performance represented by red shading ($R^2 \sim 0.8$). On the other hand, the model performs <u>generallyslightly</u>-poorer in northern counties, consistent with the results of Schauberger et al. (2017b) -where regional colder and wetter climatology <u>seems to reducereduces</u>

- 265 soybean yield sensitivity to <u>climatic fluctuationshot-dry conditions</u>. Individual diagnostic tests for models built at the county scale shows that autocorrelation and heteroscedasticity did not occur for the majority of individual models whereas model residuals are mostly normally distributed. The RESET test shows that most models are properly specified meaning that considering quadratic variables would not have improved the model fit. <u>Although quadratic associations between crop yields and climatic variables are well established</u>, these often are highlighted for seasonally averaged temperature and moisture
- 270 <u>conditions (Ray et al., 2015). Dividing the growing season into smaller periods in this study likely made these non-linear associations less relevant.</u> Finally, the VIF value is strictly smaller than 3 for almost all<u>the majority of considered models and variables showing reflecting minimal low multicollinearity concerns (Fig. A3A4).</u>

3.2 Spatial variability of model coefficients

The spatial variability of crop yield sensitivities to the selected predictors

275 The coefficient distribution for all variables, summarized across the spatial domain, is shown in Fig. 4a. Wide boxplot ranges reflect large spatial heterogeneity in coefficient estimation. This spatial variability is depicted in Fig. 4b4a-f. It shows showing county-based standardized model_coefficients and associated patterns across the spatial domain for both moisture and temperature related variables and for early, mid, and late season. Specifically selected predictors and associated timing within the season per county are shown in Fig. SA12 & SA23.

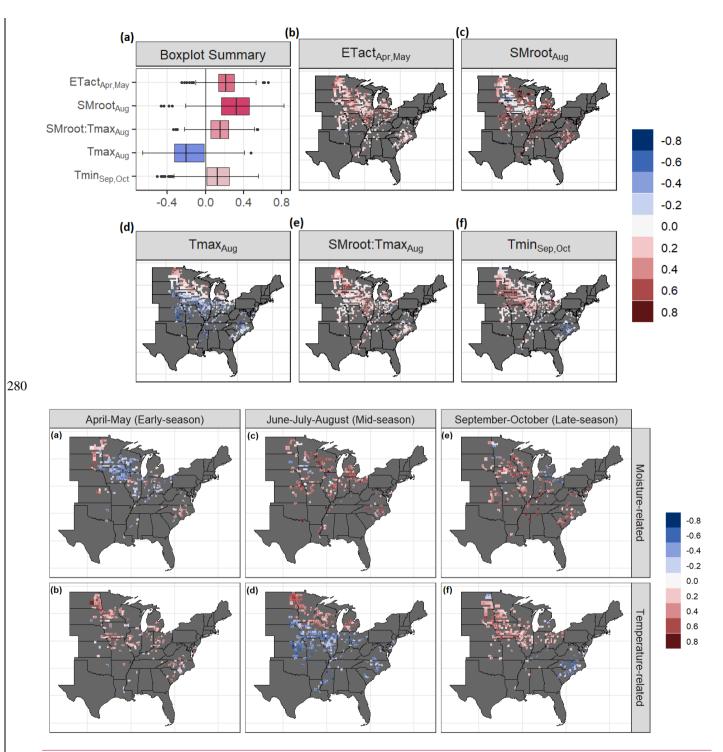


Figure 4. (a) Summary coefficient distributions across counties. The band inside the box represents the median, whereas the box depicts the 25th and 75th percentile values. The whiskers represent the maximum and minimum values as long as these are within the 1.5 interquartile range from the median. Outliers outside this range are depicted as points. (b-f) Region- and season-specific

285 estimated sensitivity coefficients for soybean yield and selected predictors. Stippling indicates statistical significance from a t-test at 95% confidence level. Values of coefficients are interpreted as the change in soybean yield standard deviation from a one-standard deviation change in the considered independent variable. In the case of interacting variables, this interpretation only applies when the other interacting variable is equal to zero.

Early season reports mainly a negative relationship between yield and moisture variables (Fig. 4a) across the majority of the

- 290 spatial domain in line with Ortiz-Bobea et al., (2019). The most frequently selected predictor is number of days with precipitation above 20 mm used as a proxy for excessive rain (Fig. 3a, A1-& SX). The signal is particularly strong and significant near Iowa and Minnesota where soils Crop sensitivity to maximum August temperatures shows a north south gradient with strongly negative regression coefficients over the southern states and sensitivities close to zero in northern states (Fig. 4d). The climatological August maximum temperature in southern states is around 30°C highlighting negative
- 295 temperature influences on yield when this value is exceeded (Fig. A4). The signal is less significant and leaning towards positive values in colder regions north of Iowa and Illinois. August root zone soil moisture is strongly positively associated to yield, with the exception of a small region near Iowa and Minnesota (Fig. 4c). Summer climatology is particularly humid over this region (Fig. A4) whereas the soil is are generally poorly drained (i.e. high clay fraction, low saturated hydraulic conductivity) (Li et al., 2019). The temperature related variable in early season (Fig. 4b) shows a positive relationship with
- 300 yields, and this can reflect both minimum and maximum temperature (Fig. A1). During the mid-season, temperature-related variables negatively affect soybean yields across the spatial domain. Exceptions are for northern states (north of Iowa and Illinois) where the sensitivity is reversed and higher temperature lead to positive effects on yield (Fig. 4d). The selected variable for the negative sensitivity (for southern states) refers mostly to maximum temperature in August whereas the positive sensitivity (for northern states) refers mostly to minimum temperature in June and July (Fig. S2Fig. A1, -A2). Moisture related
- 305 variables have a strong positive influence on yields both in the mid and late season (Fig. 4e). In particular, selected predictors are predominantly soil moisture variables in August and September. Temperature sensitivities in the late season show mostly positive effects on yield, except for counties in south-eastern states which show strong negative sensitivities (Fig. 4f). The selected late-season temperature predictor is predominantly minimum temperature for the positive associations and September maximum temperature for the negative associations over southern states (Fig. A1, -&-A2-S). Furthermore, interaction terms
- 310 between summer soil moisture and temperature variables are included in ~10% of the considered counties across the spatial domain (Fig. A3S). These interaction effects imply that the impact of summer temperature on crop yields significantly depends on the concurrent soil moisture levels in those areas. The negative effects of high temperatures are amplified during dry conditions and alleviated during wet conditions (see Sect. 3.3). Moreover, another interaction term is picked up, albeit less pronounced, between maximum August temperature and end of season minimum temperature mostly within Iowa (Fig. A3).
- 315 This might reflect increased impacts whenever anomalously hot conditions in peak summer are followed by anomalously cold conditions in September-October. The abrupt change in temperature conditions further stresses crops and reduces the potential positive effects of crop temperature acclimation -(Butler and Huybers, 2013; Carter et al., 2016).
 Such combination can make crops sensitive to the detrimental effects of excessive water on yields which could explain the negative soil moisture sensitivities here. The interaction between August heat and soil moisture variables is positive across the

- 320 majority of counties (Fig. 4e). This implies that the impact of heat in August depends on the soil moisture value. The negative effects of high temperatures are amplified during dry conditions and alleviated during wet conditions. High evapotranspiration in the early season is positively associated to yield across the spatial domain with particular strong association in central and northern states (Fig. 4b), in line with Schauberger et al. (2017b). End of season minimum temperature reveals a north south gradient in parameter estimation with significant positive effects over the colder northern regions and weaker association over
- 325 the south (Fig. 4f). The only exception is noted for south eastern states where strong negative association between yield and end of season minimum temperature is shown. Interestingly, the link between August maximum temperature and yield for those same counties is weak/not significant, suggesting that crops in this area might be reaching the temperature vulnerable stage later during the season.

3.4-3 Compound hot-dry and associated impacts

- 330 Our results show that soybean production in Southern regions is particularly sensitive to the co-occurrence of high August/September maximum Temperature and low August/September soil moisture variables (Fig. 4). The co-occurrence of low soil-moisture (5th percentile) and high temperature conditions (95th percentile) triggers the largest crop failures estimated at -2 standard deviations (calculated using domainspatially averaged model coefficients for August temperature, soil moisture and the interaction term). Extreme August hot-dry conditions (i.e. simultaneously exceeding the 95th and 5th percentiles of
- 335 temperature and soil moisture, respectively) leads to 4 times more crop yield impacts compared to extreme hot conditions alone (i.e. 95th and 50th percentiles of temperature and soil moisture, respectively) and 3 times more impacts compared to extreme dry conditions alone (i.e. 50th and 5th percentiles of temperature and soil moisture, respectively). These results are qualitatively similar when we replaced August with September soil moisture. To further illustrate the implication of including interaction terms, we focus in what follows on Illinois in what follows. Illinois is the largest soybean producing region in the
- 340 US and includes a large ratio of counties where summer moisture and temperature interactions are included in locally specified models (Fig. A3bS).

Our results show that Illinois is particularly sensitive to hot and dry conditions in August (Fig. 4e), and that therefore models perform best over this area (average \mathbb{R}^2 of 0.6 0.7). Moreover, Illinois is the largest soybean producing region in the US and hence we focus here in detail on the compounding hot dry effects in August. Figure 5a shows pooled yield observations for

345 Illinois (points) together with model predictions (contour lines) for various values of August root zone soil moisture (vertical-axis) and August maximum temperature (horizontal-axis). <u>Qualitatively similar results are obtained when we replaced August with September root zone soil moisture.</u> The coefficients for the sensitivity of soybean yields to August hot-dry conditions in Fig. 5^a are obtained from averaging all regression coefficients (i.e. for August Temperature, soil moisture and the interaction term) from all county-specific models within Illinois (i.e. 51 individual models/counties).

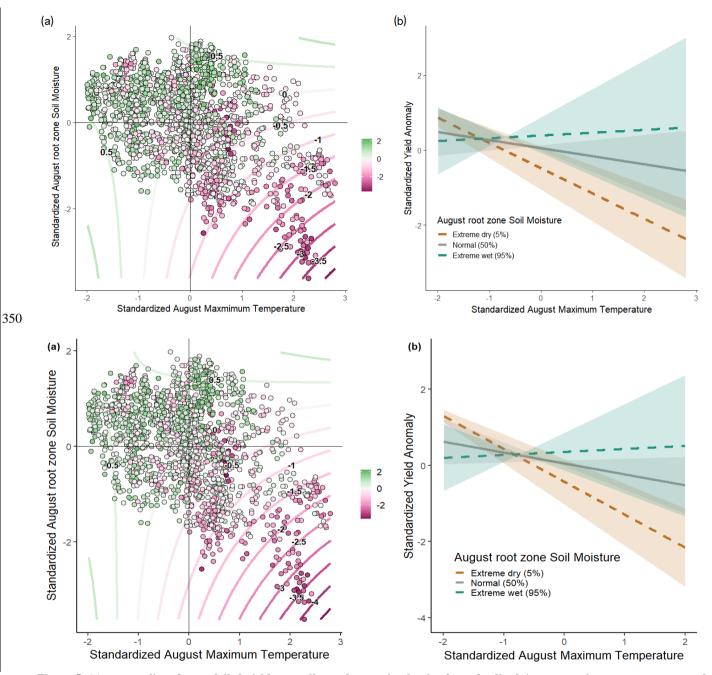


Figure 5. (a) contour lines for modelled yield anomalies under varying levels of standardized August maximum temperature and root zone soil moisture in Illinois state. Points represent observed yield values. The colour scale to the right is in the units of standardized yield<u>anomaly</u>. (b) Sensitivity of Illinois US yield anomaly to temperature change for three different root zone soil moisture percentiles (5th, 50th, 95th).

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Yield is shown to decrease for increasing hot-dry conditions both in observations and model predictions. In particular, the bottom-right_-corner (representing August temperature and soil moisture values respectively above and below the 50th

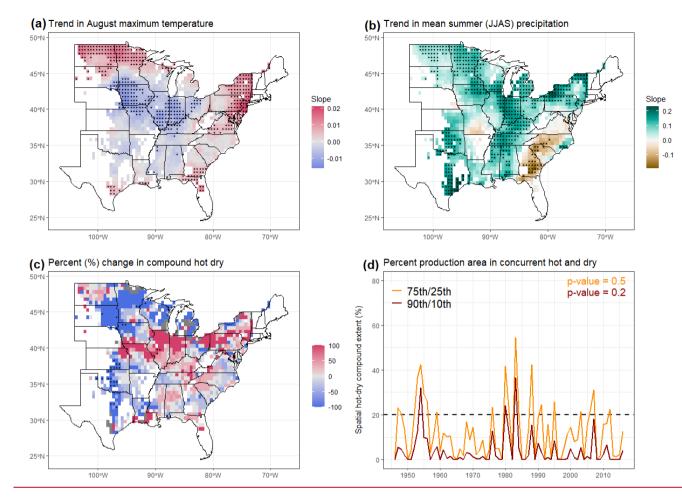
percentile) contains 75% of all observed low yields (defined as below one standard deviation).-_By including <u>interaction terms</u>, <u>LOOCV-1 model performance improved by 17% for Illinois. In particularthe interaction term</u>, we estimate that the

compounding impact of hot-dry conditions (i.e. 95th and 5th percentiles of temperature and soil moisture, respectively) in 360 August leads to an additional crop-loss of 0.65 standard deviations as compared to a model that includes all selected predictors but no interaction termsexcluding such interaction. On the other hand, the effects of extreme hot-wet conditions (95th percentile for both temperature and soil moisture values) leads to a 0.5 standard deviation positive increase in crop yield estimates when including the-interaction terms. This non-linearity is visualized in Fig. 5b showing model-derived yield sensitivities to temperature for different levels of root zone soil moisture (i.e. 5th, 50th and 95th percentiles). The association between yield and 365 August maximum temperature is strongly negative for extremely dry conditions (brown dashed line) and slightly positive for extremely wet conditions (blue dashed line). This highlights the importance of accounting for interaction effects when estimating compound impacts on crops. Yield response to hot-wet conditions is nevertheless subject to high uncertainty (see shaded uncertainty range in Fig. 5b) as these conditions do not occur often and are represented by few observations (upper-370 right corner in Fig. 5a). The rarity of these events is expected owning to the negative correlation between moisture and temperature over summer (Zscheischler and Seneviratne, 2017)-. It follows that wet conditions generally limit exposure rather than sensitivity to very high temperature. Still, temperature sensitivities during wet conditions are significantly different from

those during dry conditions (Fig. 5b).

3.54 Changes in compound hot-dry from 1946 to 2016

Linear trends for summer precipitation (JJAS) over the period 1946 to 2016 show significant increases particularly over the Midwest region (Fig. 6b). Only south-eastern states show significant drying trends. Maximum August temperature trends show significant cooling over the Midwest region but warming for north-eastern, north-western and southern states (Fig. 6a). Moreover, early and late season minimum temperature trends indicate warmer conditions across the spatial domain whereas early season number of wet days trend indicates wetter conditions in spring- (see Fig. A55). Though summers generally got wetter and cooler in the eastern part of the Midwest and north eastern US regions, the percent-change in the number of concurrent hot and dry August-summer months (i.e. 90th and 10th percentiles of August maximum temperature and summer precipitation, respectively) between 1946-1980 and 1982-2016 shows an increase in frequency here (Fig. 6c). This might have implications as compound hot-dry events appear to have increased in frequency in high producing regions, despite the apparent cooling and wetting patterns identified by univariate trends.



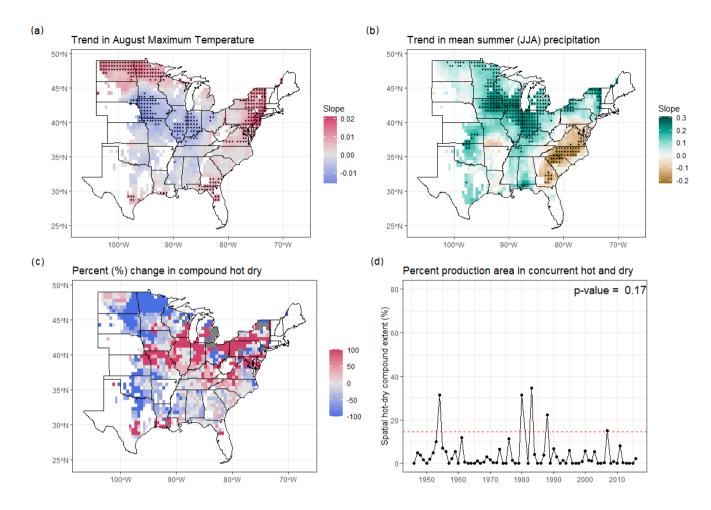


Figure 6. (a) Linear regression slope of August maximum temperature. (b) Linear regression slope for summer (JJAS) precipitation. (c) Percent (%) change in concurrent dry (Summer JJAS precipitation < 10th percentile) and hot (August Maximum Temperature > 90th percentile) during 1982–2016 relative to 1946–1980. (d) Time-series of percent producing regions in hot and dry conditions.
Trends in (a, b and d) are calculated for the period 1946 to 2016. Stippling in (a), <u>-and</u> (b) <u>and (c)</u> indicates statistical significance at the 955% confidence level. P-value in d) corresponds to the—Mann–Kendall monotonic trend test. <u>Red-Black</u> dashed line in (d) represents a 15% threshold marking years with a large (>15%) spatial hot-dry extent.

Time series of percent production area in concurrent hot and dry conditions reflects the spatial extent of such conditions over the years (Fig. 6d). The red-black dashed line represents a threshold set at 1520% exceeded by a number of years_-(i.e. 1947, 1948, 1953, 1954, 1955, 1956, 1959, 1976, 1980, 1983, 1984, 1988, 1991, 1995, 2003, 2006, 2007, 20121956, 1980, 1983, 1988 and 2007 when using the 75th-/25th percentile hot-dry time series). More than 60% of those years coincide with La Niña like conditions_All those years except 1980 coincided with a developing La Niña summer reported, which have been shown to impact US crop production (Anderson et al., 2019; Iizumi and Sakai, 2020). FurthermoreMoreover, we note a high frequency of large-scale hot-dry events in specific periods such as the 1950s and 1980s. These segmented periods of high intensity events
suggest the potential important role of decadal climate variability in the occurrence of hot-dry conditions. These can be related

to Llow-frequency sea surface temperature variations in the pacificAtlantic and AtlanticPacific oceanssuch as the pacific decadal oscillation (PDO) have been indeed shown to have an influence on local precipitation and temperature levels over eastern US (Vijverberg et al., 2020). -A large fraction of the production area under such hot-dry conditions create risks for country level agricultural production as regions are no longer able to balance out losses at the local scale. Here again, despite the dominant cooling and wetting trends over the US (Fig. 6-a & d), no significant up- or downward trend was found in the fraction of US under hot-dry conditions over time for both the 75th/25th and 90th/10th time-series.

3.5 Temperature and moisture couplings during summer hot-dry events

To better understand why compound hot-dry conditions have not changed, despite significant trends towards wetter summers and cooler August maximum temperatures, we analyse local land-atmosphere couplings. It has been hypothesized that during

410 dry conditions, the actual evapotranspiration reduces, cancelling the land-change induced cooling effect and prompting a return to historic high temperature extremes (Mueller et al., 2016). Monthly correlations between root-zone soil-moisture (SMroot), maximum temperature (Tmax) and actual evapotranspiration (AET) are used to estimate the coupling strength during hot-dry summers and normal summers. The subset of hot-dry events in this case is constructed from years when more than 20% of the total harvested area is under hot-dry conditions, defined using the 75th and 25th percentiles of August maximum temperature and summer precipitation (JJAS) respectively (i.e. years when the orange line is above the dashed black line in Fig. 6d).

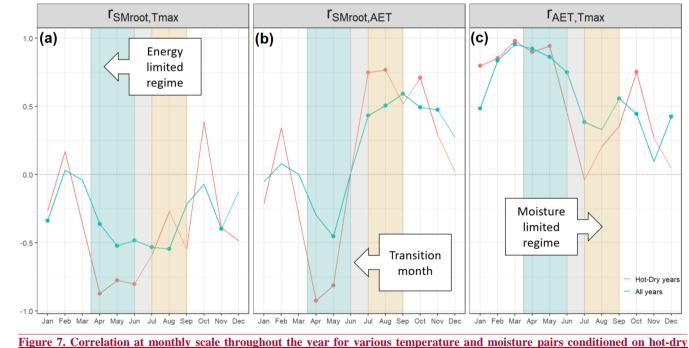


Figure 7. Correlation at monthly scale throughout the year for various temperature and moisture pairs conditioned on hot-dry events. Dots indicate statistical significance at the 95% confidence. Shaded regions represent important differences in the couplings that can play a critical role in the development of hot-dry events.

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We observe that summer hot-dry years are characterized by a stronger negative coupling between soil moisture and temperature during spring (April-May) when compared to a typical year (Fig. 7a). We interpret this negative coupling as indicative of warmer and drier springs. These conditions create a stronger negative coupling between evapotranspiration and soil moisture as evapotranspiration rates are enhanced by warmer temperatures, in turn, rapidly depleting soil moisture reserves (Fig. 7b).

- 425 The timing when the coupling between evapotranspiration and soil moisture sign shifts reflects a critical moment in the system when soil moisture becomes limiting. We observe that this regime-shift is much more pronounced during hot-dry years (i.e. stronger negative coupling in April-May and stronger positive coupling in July-August) (Fig. 7b). June is a transition month. The moment of the regime shift (around June) coincides with the ceasing of the spring coupling between evapotranspiration and temperature during hot-dry years (Fig. 7c). We interpret this ceasing of the coupling between evapotranspiration and
- 430 <u>maximum temperature as an indicator of total depletion of moisture in the soils, and thus extra energy (via higher temperatures)</u> cannot lead to more evaporation. We are thus in a moisture-limited land-atmosphere coupling regime. During normal years, still significant coupling between evapotranspiration and maximum temperature exists in July-September indicating that the soils are not fully depleted. Spatially, the ceasing of the land-surface induced cooling effect is present over most of the soybean harvesting region going from June to September for hot-dry years (Fig. A6). In summary To summarize, we show that summer
- 435 <u>hot-dry events are associated towith warmer and drier springs. These conditions -springs that lead tofavour faster and more intense depletion of soil moisture-depletion. Dry soils limit the evaporative cooling effect as captured by the annulled co-variability between actual evapotranspiration and temperature leading to amplified hot and dry conditions in summer (Fig. 7c). T-amplifying hot and dry conditions over summer. The ceasing of the land surface induced cooling mechanism appears to be a critical element in the development of hot dry events. This provides evidence in support of the initial hypothesis that highlights the important role of land-atmosphere feedbacks in explaining the absence of a trend in summer hot-dry events.</u>
- despite summer wetting and cooling trends over the soybean production region in the US.

4 Discussion

Predictors here are determined statistically, nevertheless, we aimed for a <u>unique_restricted_set of_-moisture and temperature</u>
variables for all US counties to facilitate the physical interpretation of climatic drivers affecting soybean yield variability. This is in line with other studies that constructed semi-empirical crop models at the grid-cell level-relying on a statistical framework driven by well-known physiological variables (Ben-Ari et al., 2018; Gornott and Wechsung, 2016; Schauberger et al., 2017b). The frugal method approach we used to select predictors means leaving outimplies that potentially useful and physiologically-relevant variables such as radiation and excessive precipitationvapour pressure deficit are omitted. Although their effects can be implicitly accounted for in the temperature and moisture variables used, --light exposure, for instance, certainly plays a key

role in crop productivity (Farquhar et al., 2001; Rigden et al., 2020). Nevertheless, t Thishe choice is made as the least-squares model fit is highly sensitive to the ratio of predictors to the number of observations (James et al., 2013). Ideally, crop-

observations (35 here) should be much larger than the number of predictors to avoid <u>the risk of</u> overfitting. Furthermore, including highly-correlated predictor variables (e.g. radiation and temperature) affect model parameter <u>estimation</u><u>estimation</u>

455 and complicates physical interpretation of drivers. <u>A reduced set of predictor variables where shared information between variables is minimized provides an easily interpretable and robust model for assessing sensitivity of soybean crops to climate and weather variability (Ben Ari et al., 2018; Gornott and Wechsung, 2016; Lobell and Burke, 2010; Schauberger et al., 2017b). Future studies can disentangle these mechanisms for a more detailed data-driven assessment of climate and crop yield sensitivities. It is also possible to use more complex machine learning models such as random forests although these often tend</u>

to obscure result interpretation and do not always <u>yield-lead to</u> better predictions (Vogel et al., 2019, 2021). Note that nonclimatic seasonal influences on crop yields are ignored in this study. These include planting densities, sowing dates, fertilizer applications and other socio-economic factors. This simplification is done as spatially-explicit time series for such components are rare and difficult to obtain (Schauberger et al., 2017b). Some of these factors were shown not to necessarily improve model performance in a case study done on crop yields in Germany (Gornott and Wechsung, 2016). Nevertheless, future studies should include these in whenever this becomes possible for extended time periods as climate has been shown to influence seasonal management practices for farmers in the US (Carter et al., 2018b).

470 We found that soybean yields were predominantly driven by heat and drought conditions occurring during the vulnerable summer crop reproductive stage. In particular, August and September months was were highlighted as key months for soybean production in line with results from previous studies (Mourtzinis et al., 2015; Ortiz-Bobea et al., 2019; Zipper et al., 2016). Furthermore, we noted a significant interaction effect between August summer maximum temperature and soil moisture variables modulating the final impact on yield. -Drought and heat induce different growth inhibition patterns that can act 475 simultaneously to reduce crop photosynthetic rates and eventual yield levels (Suzuki et al., 2014). August mean maximum temperature was found to be negatively associated with soybean yields for values exceeding 30°C (i.e. average August maximum temperature value for a large part of the considered counties). This is -in line with other studies reporting-that reported non-linear association between soybean and temperature where the relationship is mildly positive up until the 30°C mark and then declines sharply due to heat stress (Schauberger et al., 2017a; Schlenker and Roberts, 2009). 480 Nevertheless Moreover, here we found that this relationship was dependent on concurrent soil moisture conditions where wet soils dampen the negative effect of high temperatures on yield via evaporative cooling. This-A result is also in line withsupported by previous studies reporting the decoupling effect of irrigation on the relationship between heat stress and yield (Carter et al., 2016; Schauberger et al., 2017a; Schlenker and Roberts, 2009; Siebert et al., 2017; Troy et al., 2015). On the other hand, low moisture levels induce stomatal closure which leads to reduced latent heat flux and an increase in canopy 485 temperature well above atmospheric temperatures increasing the crop sensitivity to hot conditions (Carter et al., 2016; Siebert et al., 2017). Such dependency highlights the important need to account for both variables simultaneously when assessing their

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<u>impacts on crop yield variability</u> Such interaction, although well documented in the literature on crop physiological response to primary abiotic stressors is rarely considered in large scale statistical analyses of climate impact on crop yields. This suggests a potential overestimation of temperature effects during wet conditions and an underestimation of compound hot dry impacts

- 490 in previous reports (Carter et al., 2018a; Leng et al., 2016; Siebert et al., 2017; Suzuki et al., 2014).-.-Our analysis further highlighted reported early season evapotranspiration excessive precipitation and minimum and maximum temperature conditions in addition to late season minimum temperature as important drivers of soybean yield variability. Early season excessive precipitation sensitivity likely reflects damaging plant field establishment conditions related to restricted root development, nutrient leaching and disease susceptibility (Li et al., 2019; Ortiz-Bobea et al., 2019). -High evapotranspiration
- 495 <u>minimum and maximum temperature</u> in the early season positively associated to yield reflects mainly non limiting energy conditions as moisture levels are expected to not be restrictive in early spring. This can imply both a reduced frost risk in addition to a potentially longer growing season where soybean yield potential is maximized (Bastidas et al., 2008; Mourtzinis et al., 2019). End of season frost has also been reported as an important risk factor for soybean crops particularly in the northern states, and we interpret the predictor of minimum temperature during September and October as reflective of such conditions.
 500 -These identified drivers of impact can serve as a basis for effective early warning systems that provide valuable information to decision makers (Merz et al., 2020). Acting in advance can be critical to avoid crop loss and associated socio-economic consequences. For instance, a short period of drought during the reproductive stage is reported to cause non-reversible damage to soybean yields (Daryanto et al., 2017). Hot and dry conditions in eastern US over summer has been shown to be forecastable at long lead times (~50 days ahead), associated with sea surface temperature anomalies over the northern Pacific Ocean (McKinnon et al., 2016; Vijverberg et al., 2020). Future work can further explore the link between drivers of compound hazards

impacting yields to facilitate the development of actionable tools for stakeholders.

We showed that historic changes in climate have not increased the overall climate risk for rainfed soybean production in the US. This is in line with other studies that looked at the contribution of historic climate trends on soybean and maize yields in the US (Butler et al., 2018; Ray et al., 2019). This is particularly the case in the most northern states where the occurrence of compound hot-dry events has mostly decreased (Fig. 6d). Interestingly, soybean cropping regions have also shifted northwesterly in the US taking advantage of such changes in climate (Sloat et al., 2020). Increasing trend in number of wet days during spring can lead to detrimental change for rainfed soybean production. Nevertheless, Lesk et al., (2020) recently highlighted that the association between heavy rainfall and US crop yields can be different and more complex when studied at

- 515 <u>sub-daily resolution emphasizing that further investigation is needed in that regards.</u> -The summertime cooling is a welldocumented phenomenon over US agricultural regions (Nikiel and Eltahir, 2019) and is likely attributable to agricultural intensification in the region <u>although among</u> other driving processes such as decadal variability and the role of aerosol emissions also play a role (Alter et al., 2018; Lesk and Anderson, 2021; Mueller et al., 2016; Nikiel and Eltahir, 2019). With respect to the role of agriculture, <u>A</u> a higher density of crops supported by increasing fertilizer rates leads to higher
- 520 evapotranspiration rates which in turn induce large scale evaporative cooling and contribute to increasing precipitation (Basso

et al., 2021; Mueller et al., 2016). Nevertheless, we highlighted that in key producing regions like Illinois, compound hot-dry events seem to have increased in frequency recently, despite the absence of a <u>summer-mean</u> drying or warming trend. Potentially, during dry conditions, the actual evapotranspiration reduces, cancelling the land-change induced cooling effect and prompting a return to historic high temperature extremes (Mueller et al., 2016). We illustrated this mechanism by analysing

- 525 the evolution of land-atmosphere coupling within the growing season, captured by monthly correlations between actual evapotranspiration, root zone soil-moisture and maximum temperature. We interpreted positive correlation values between actual evapotranspiration and maximum temperature as indicative of a general land-surface induced cooling effect. During hot-dry years, this evaporative cooling ceased at the onset of summer months.- We showed that this was associated to stronger negative coupling between evapotranspiration-soil moisture and soil moisture-temperature in spring. Such conditions lead to
- 530 fast soil moisture depletion and favour a moisture limited regime that amplifies extreme summer hot-dry conditions and associated soybean impacts (Sippel et al., 2016). Although we showed that warmer and drier springs lead to higher yields, potentially connected hot and dry summer conditions lead to disproportionately negative impacts on final crop yields. -Future risk assessments should account for such non-linear effects. Over the Midwest US, climate models project warmer summers which is likely to enhance the coupling between moisture and temperature via land-atmosphere feedbacks leading to a likely possible increase in the amplitude and frequency of compound hot-dry conditions (Cheng et al., 2019; Zscheischler and Seneviratne, 2017). Although annual precipitation levels are expected to remain constant or even increase, climate models generally project increased dry day length and decreased summer soil moisture levels (Dai, 2013; Dirmeyer et al., 2013;
- Wuebbles et al., 2014a, 2014b). Future research should quantify whether such trends could lead to an increase of hot-dry August-summer months in the future. Nevertheless, high uncertainty remains with respect to atmospheric dynamical changes
 including quasi-stationary Rossby waves which are a key driver of hot-dry conditions in the eastern US as well as other midlatitude regions (Di Capua et al., 2020; Coumou et al., 2014; Kornhuber et al., 2019; Shepherd, 2014; Winter et al., 2015). Until such contradictions are resolved, future impacts of climate change on US agricultural production remain uncertain. The Storylinestoryline approach has been proposed as an important tool to illustrate such epistemic uncertainty and can be explored in future studies with important consequences on current and future policy and decision making (Shepherd, 2019).

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Here we focused on local types of compound events, however, global food supply is highly dependent on production in various countries. Spatially compounding events will be important to study in future assessments in order to understand large scale risk associated to breadbasket failures. Here we qualitatively identified that <u>most-a considerable number</u> of the large extent hot-dry conditions occurring over the US are <u>associated-coinciding with La Niña like conditions to ENSO teleconnections</u>. These are also highly influential over the South American continent where soybean production including the US account for more than 80% of total global supply (Anderson et al., 2017; Iizumi and Sakai, 2020; Wellesley et al., 2017). Other examples of teleconnections are mid-latitude Rossby waves, particularly wave number 5, which has phase-locking behaviour in the northern hemisphere mid-latitudes driving simultaneous summer positive temperature anomalies over Midwest US, eastern

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Europe, and east Asia (Kornhuber et al., 2019). This is particularly of concern to soybean production when taking into

555 consideration upcoming soybean hotspot production regions such as Russia and Ukraine (Deppermann et al., 2018).

5 Conclusion

We presented a simple statistical framework that can identify climatic variables influencing soybean yield variability in the US at specific moments within the growing season. We found that compound August summer hot-dry conditions lead to the largest impacts on yield, i.e. beyond the estimated additive effects of each stressor separately. Furthermore, we identified early-560 season evapotranspiration minimum and maximum temperature in addition to precipitation, and late-season minimum temperature and soil moisture to be important factors affecting soybean yield in the US. Understanding of these seasonally dependent crop-sensitivities paves the way for more effective early-warning tools that target timely drivers of yield variability throughout the growing season. The long-term cooling and wetting trend in summer, over large areas of our domain, has generally been beneficial for soybean. Nevertheless, we showed that the frequency of extreme hot-dry conditions remained 565 largely unchanged over the full region, and increased in a key region like Illinois where crops are especially sensitive to such extremes. Furthermore, we showed that hot-dry events are characterised by stronger negative spring coupling between evapotranspiration-soil moisture and soil moisture-temperature leading to fast soil moisture depletion in spring and a reversal in the land-surface cooling mechanism over summer prompting important soybean yield impacts. Given that climate models project summer warming and general declines in soil-moisture (albeit with substantial uncertainty) for the Midwest, crop 570 sensitivities to compound hot-dry extremes are likely to present important future risks for US soybean production.

Appendix A: Additional tables and figures

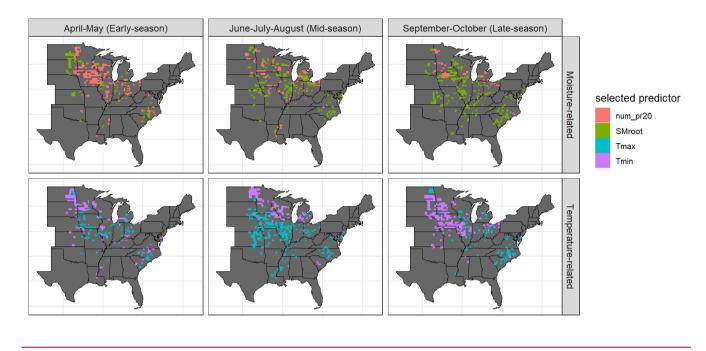
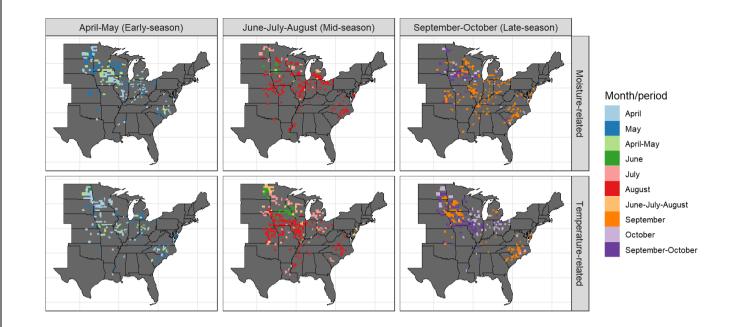


Figure A 1. Selected predictors per county based on the full dataset



575 Figure A 2. Selected timing of predictors per county based on the full dataset

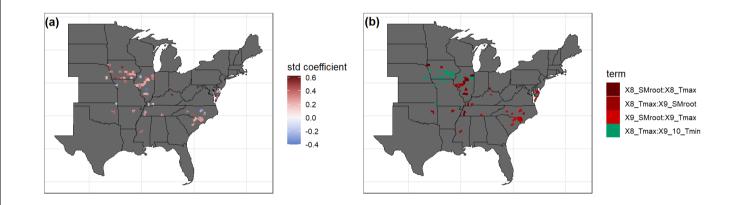


Figure A 3. (a). Standardized coefficients for interaction terms per county selected based on the full dataset. (b). Type of interactions selected per county based on the full dataset.

580 Table A 1. List of all considered variables tested individually as potential predictors for the statistical model. The selected heat and moisture variables per period are highlighted in red.

Best heat and moisture related variables for each univariate model based on BIC

Spring		St.	ımmer		Fall			
Name	BIC	Туре	Name	BIC	Туре	Name	BIC	Type
X4_5_ETact	38101.02	moisture	X8_SMroot	35671.27	moisture	X9_SMroot	35478.40	moisture
X4_5_Tmin	38114.13	heat	X6_7_8_ETact	36494.53	moisture	X9_ETact	35497.50	moisture
X4_ETact	38126.94	moisture	X8_Tmax	36502.21	heat	X9_10_ETact	35644.31	moisture
X4_Tmax	38135.90	heat	X8_rsds	36667.57	heat	X9_10_SMroot	35782.78	moisture
X4_5_Tmax	38158.38	heat	X6_7_8_precip_avg	36672.72	moisture	X10_SMroot	36388.79	moisture
X4_SMroot	38163.37	moisture	X8_precip_avg	36684.10	moisture	X10_ETact	37251.21	moisture
X5_ETact	38165.18	moisture	X6_7_8_num_wet	36719.33	moisture	X9_10_Tmin	37791.13	heat
X4_Tmin	38175.23	heat	X6_7_8_SMroot	36883.95	moisture	X10_Tmin	37940.51	heat
X5_Tmin	38176.42	heat	X8_ETact	36981.19	moisture	X9_Tmin	38003.55	heat
X4_5_SMroot	38181.28	moisture	X8_num_tx30	36986.02	heat	X9_10_rsds	38062.98	heat
X5_SMroot	38213.79	moisture	X8_num_wet	37065.40	moisture	X9_rsds	38095.34	heat
X5_num_pr20	38218.38	moisture	X6_7_8_rsds	37232.01	heat	X9_num_wet	38159.20	moistur
X4_5_num_pr20	38225.77	moisture	X7_SMroot	37321.64	moisture	X9_10_num_tx30	38179.84	heat
X5_num_pr30	38236.84	moisture	X6_7_8_Tmax	37387.85	heat	X9_10_num_wet	38180.22	moistur
X4_5_num_pr30	38239.45	moisture	X8_num_pr20	37397.33	moisture	X9_num_tx30	38193.27	heat
X4_5_precip_avg	38243.69	moisture	X6_7_8_num_pr20	37413.15	moisture	X10_rsds	38211.16	heat
X5_precip_avg	38245.36	moisture	X7_ETact	37457.28	moisture	X10_Tmax	<u>38219.46</u>	heat
X5_Tmax	38253.65	heat	X7_precip_avg	37506.25	moisture	X9_10_precip_avg	38225.60	moistur
X4_num_tx30	38272.67	heat	X6_7_8_num_tx30	37556.24	heat	X10_num_tx30	38237.57	heat
X4_precip_avg	38273.35	moisture	X7_num_wet	37592.68	moisture	X9_precip_avg	38241.65	moistur
X4_num_wet	38274.49	moisture	X7_rsds	37717.13	heat	X9_Tmax	38247.62	heat
X4_rsds	38275.85	heat	X8_num_pr30	37747.78	moisture	X9_10_num_pr20	38248.19	moistur
X4_5_num_tx30	<u>38276.22</u>	heat	X7_Tmax	37757.88	heat	X9_num_pr20	38263.08	moistur
X4_num_pr20	38276.54	moisture	X6_7_8_num_pr30	37807.09	moisture	X9_10_num_pr30	38271.85	moistur
X5_num_tx30	38280.31	heat	X6_ETact	37846.45	moisture	X10_num_pr30	38274.50	moistur
X4_num_pr30	38281.21	moisture	X7_num_pr20	37913.19	moisture	X10_precip_avg	38280.89	moistur
X5_num_wet	38288.94	moisture	X6_SMroot	37925.90	moisture	X10_num_wet	38281.09	moistur
X4_5_num_wet	38289.06	moisture	X7_num_tx30	37996.85	heat	X10_num_pr20	38281.21	moistur
X4_5_rsds	38289.74	heat	X8_Tmin	38054.98	heat	X9_10_Tmax	38284.94	heat
X5_rsds	38290.43	heat	X7_num_pr30	38067.72	moisture	X9_num_pr30	38286.41	moistur
			X6_num_wet	38163.44	moisture			
			X6_Tmin	38186.92	heat			
			X6_rsds	38221.41	heat			
			X6_precip_avg	38239.57	moisture			

X6_num_tx30	38240.49	heat
X6_7_8_Tmin	38260.96	heat
X6_num_pr20	38263.44	moisture
X6_Tmax	<u>38281.99</u>	heat
X7_Tmin	38282.28	heat
X6_num_pr30	38289.71	moisture

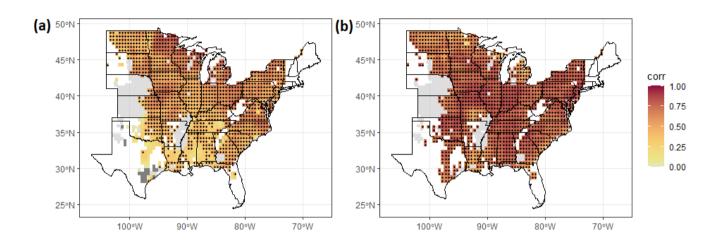


Figure A 1. (a) Correlation plot between CRU April-May minimum temperature<u>wet days</u> and GLEAM <u>W5E5</u> April-May
 585 actual evapotranspiration<u>number of days with precipitation above 20mm</u> over the period of study (1982-2016). (b) Correlation

plot between CRU JJAS precipitation average and GLEAM August <u>September average</u> root zone soil Moisture over the period of study (1982-2016). Stippling indicates statistical significance at the 95% confidence level.

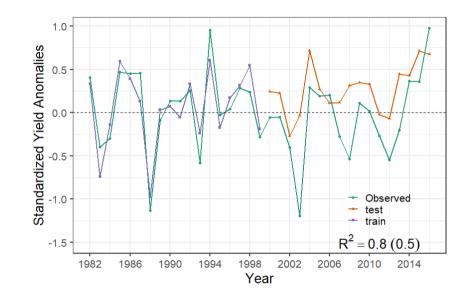
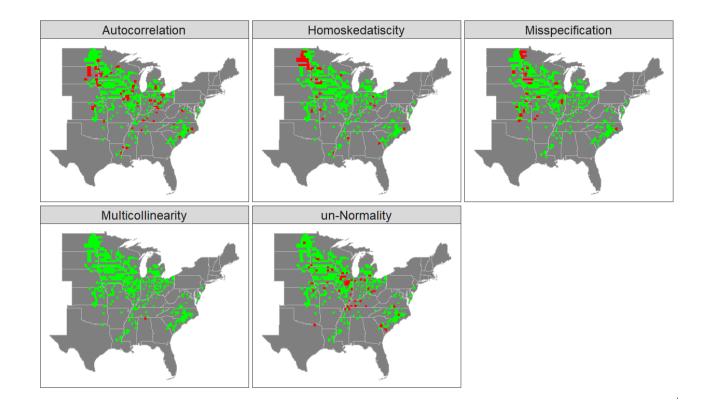


Figure A 2. Train test split validation approach where model is trained over 50% of the data (Blue line) and tested over the remaining 590 50% (Red).



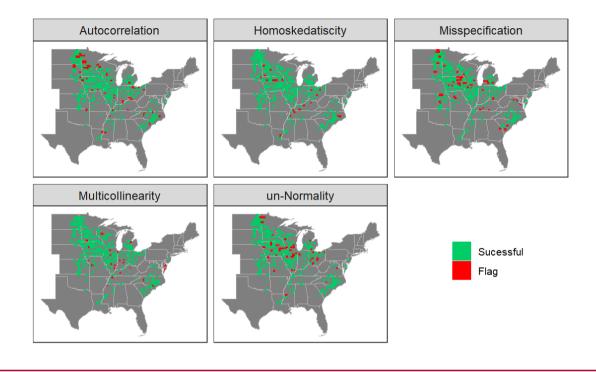


Figure A <u>34</u>. <u>Statistical Diagnostic</u> test results for the <u>USfitted models</u>. Green indicates a "successful" test, i.e. no problem, while red indicates a rejection of the respective H0 of no autocorrelation/heteroscedasticity/ misspecification/multicollinearity/un-normality. Multicollinearity is checked with the variance inflation factor and marked in red if any of the variables report a value >3.

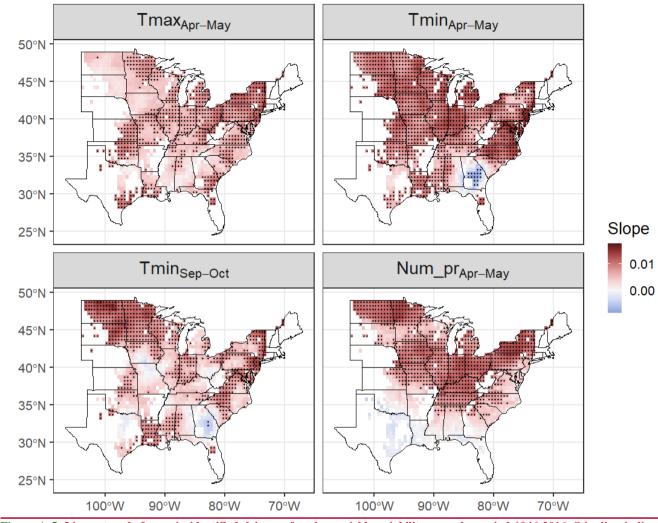
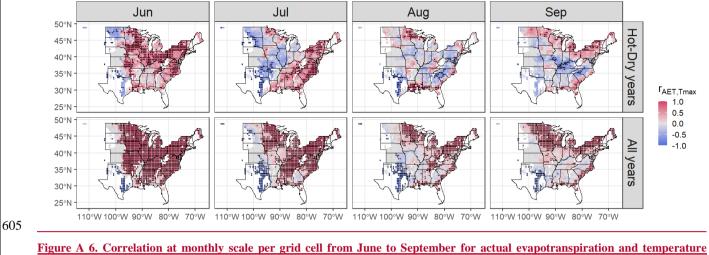


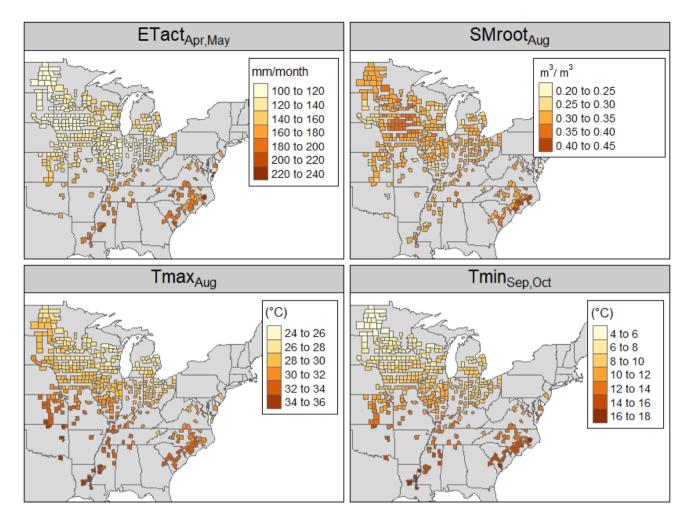
Figure A 5. Linear trends for main identified drivers of soybean yield variability over the period 1946-2016. Stippling indicates statistical significance at the 95% confidence level. Trends for moisture and temperature variables over summer are displayed in the main text.

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(a) Linear regression slope for April-May (spring) minimum temperature. (b) Linear regression slope for September-October (fall) minimum temperature. Stippling indicates statistical significance at the 95% confidence level.









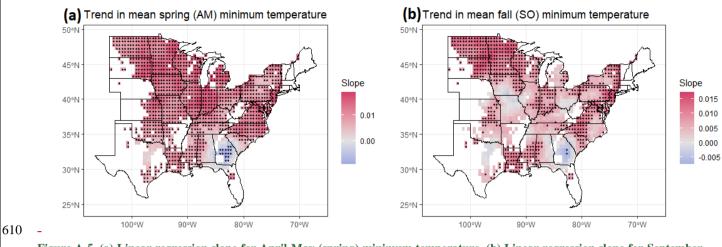


Figure A. 5. (a) Linear regression slope for April-May (spring) minimum temperature. (b) Linear regression slope for September-October (fall) minimum temperature. Stippling indicates statistical significance at the 95% confidence level.

615 *Code availability.* The code is available upon request, by contacting the corresponding author.

Data availability. Data used in this study are freely available in the cited literature.

Author contributions. RH and DC designed the study. RH performed the analysis and wrote the initial draft of the
 manuscriptpaper. All authors contributed toation the development of the analysis, the interpretation of the results and to the
 writing of the paper. All authors discussed the analysis and results, and contributed to the writing of the paper. DC revised the

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