Return Levels of Temperature Extremes in Southern Pakistan

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10 Abstract. Southern Pakistan (Sindh) is one of the hottest regions in the world and is highly vulnerable to temperature extremes. In order to improve rural and urban planning, it is useful to gather information about the recurrence of temperature extremes. In this work, return levels of the daily maximum temperature T_{max} are estimated, as well as the daily maximum wet-bulb temperature TW_{max} extremes. We adopt the Peaks over threshold (POT) method, which have not yet been used for similar studies in this region. Two main datasets are 15 analyzed: temperatures observed in nine meteorological stations in southern Pakistan from 1980 to 2013, and the ERA Interim (ECMWF re-analysis) data for the nearest corresponding locations. The analysis provides the 2, 5, 10, 25, 50 and 100-year Return Levels (RLs) of temperature extremes. The 90% quantile is found to be a suitable threshold for all stations. We find that the RLs of the observed T_{max} are above 50°C in northern stations, and above 45°C in the southern stations. The RLs of the observed TW_{max} exceed 35°C in the region, which is 20 considered as a limit of survivability. The RLs estimated from the ERA Interim data are lower by 3°C to 5°C than the RLs assessed for the nine meteorological stations. A simple bias correction applied to ERA Interim data improves the RLs remarkably, yet discrepancies are still present. The results have potential implications for the risk assessment of extreme temperatures in Sindh.

25 Key words

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Extreme temperature, return levels, peak over threshold, Generalized Pareto Distribution, declustering.

1 Introduction

Extreme maximum temperature events have received much attention in recent years, because of the associated dangerous impact on the increased risk of mortality. Additionally, climate change scenarios suggest that in most regions the probability of occurrence of extremely high temperature is very likely to increase in the future (Sheridan and Allen, 2015). An example of the potential impact of raising maximum temperatures is the recent heat wave in southern Pakistan (Sindh), which occurred between June 17th and June 24th 2015 and broke all the records with a death toll of 1400 people, and over 14000 people hospitalized. The temperatures in different cities of the Sindh region were in the range of 45°C - 49°C during the event (Imtiaz and Rehman, 2015). Karachi had the highest number of fatalities (1200 people approximately). The Pakistan Meteorological department issued a technical report stating a very high heat index (measuring the heat stress on humans due to high temperature and relative humidity) during this heat wave (Chaudhry et al., 2015).

In summer, Sindh becomes very hot and with the arrival of a monsoon the humidity increases in the region (Chaudhry and Rasul, 2004). The extremely hot and humid conditions can have lethal effects, and can impact the human habitability of a region (Pal and Eltahir 2015). The human body generally maintains the temperature

around 37°C. However, the human skin regulates at or below 35°C to release heat (Sherwood and Huber, 2010). Under high levels of the moisture content in the atmosphere, the human body cannot maintain the skin temperature below 35°C and can develop ailments like hyperthermia, heat strokes and cardiovascular problems. Hyperthermia is a condition where extremely high body temperature is reached, resulting from the inability of the body to get rid of the excess heat. It occurs mostly when temperature and relative humidity levels are extremely high at the same time. Hyperthermia can occur even in the fittest human beings, if exposed to an environment where wet-bulb temperature is greater than 35°C for at least six hours.

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This study devotes special attention to Sindh because of its exposure to the frequent and intense temperature extremes in the past (Zahid and Rasul, 2012). This region is considered as one of the most vulnerable regions in Pakistan. Sindh stretches from 23.5° N – 28.5° N and 66.5°E - 71.1°E, and is bounded on the west by the Kirthar Mountains, to the north by the Punjab plains, on the east by the Thar desert and to the south by the Arabian Sea (Indian Ocean), while in the center there is a fertile land around the Indus river. The Indus river is the source of water for the agricultural lands. Cotton, wheat and sugar cane are grown on the left bank of the Indus and rice, wheat and gram on the right bank (Chaudhry and Rasul, 2004). Cotton is the cash crop of the country.

The climate in Sindh is arid and subtropical with less than 250 mm annual rainfall. The temperature frequently exceeds 45°C in summer (May-September) and the minimum average temperature recorded during winter (December- January) is 2°C. Table 1 shows the mean monthly climatic characteristics of the region from 1980-2010. Figure 1 shows the spatial distribution of all nine weather stations of Pakistan meteorological department, and the ERA Interim grid points close to the corresponding locations. High population density, limited resources, poor infrastructure and high dependence of the local agriculture on climatic factors, mark this region as highly vulnerable to the impacts of climate change.

The Intergovernmental Panel on Climate Change (IPCC) scenarios estimates for this region an increase in the surface temperature of the order of 4°C by the end of 2100. This may significantly reduce crop yields, and cause huge economic losses to the country (Islam et al., 2009; Rasul et al., 2012; IPCC, 2012, 2014). Furthermore, the risks of heat strokes, cardiac arrest, high fever, diarrhea, cholera and vector borne diseases might increase. Heat waves became more frequent and intense during 90's in Southern Pakistan. Zahid and Rasul (2010) reports a significant rise in the heat index and heat wave events longer than ten days in Sindh. The enhanced mortality rate related to the heat waves is a serious problem, and two obvious examples are the 1991 and the previously mentioned 2015 heat waves (Imtiaz and Rehman, 2015).

The analysis of extreme climatic events is a very active area of research in geosciences (Christidis et al., 2005, 2010; Tebaldi et al., 2006; Zwiers et al., 2011; Morak et al., 2011, 2013). In order to facilitate and standardize the analysis of extremes, the World Meteorological Organization (WMO) has suggested 27 specific climate indices, like the number of hot days, cold days, wet days, dry days, etc. (Tank et al., 2006; 2009, Frisch et al., 2002; Choi et al., 2009; Lustenberger et al., 2014). The investigation and analysis of such climate indices has now reached a high level of popularity.

Extreme value theory (EVT) provides the statistical basis for increasingly widespread quantitative investigation of extremes in climate studies (Coles, 2001, Zhang et al., 2004; Brown et al., 2008; Faranda et al., 2011; Acero et al., 2014). The peaks over threshold (POT) approach aims at describing the distribution of the exceedances of the stochastic variable of interest above a threshold. Under very general conditions, the exceedances are asymptotically distributed according to the Generalized Pareto Distribution (GPD). GPD has remarkable properties of universality when the asymptotic behavior is considered (Lucarini et al., 2016), while one can expect that the threshold level above which the asymptotic behavior is achieved depends on the characteristics of the analyzed time series. In particular, when looking at spatial fields, the threshold level depends on the geographical location.

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In this study, we have chosen to use the POT method to assess the temperature extremes in the Sindh region, because it provides a more efficient use of data and has better properties of convergence when finite datasets are considered (Lucarini et al., 2016). Additionally, we are here interested in investigating the actual tails of the distributions, so the POT approach is more appropriate. It is applied for studying temperature extremes in different regions of the world (Burgueño et al., 2002; Nogaj et al., 2006; Coelho et al., 2007; Ghil et al., 2011). However, to our knowledge, the POT method was never used to analyze the risk of temperature extremes in Sindh. The POT approach provides estimates of return periods and the return levels also for time ranges even longer than what is currently observed. This information and this predictive power can be beneficial for policy makers and other stakeholders. Note that this is exactly the kind of information planners need when, e.g., designing infrastructures that are deemed to last a very long time.

It is useful to consider two indicators of extremely hot conditions: (1) temperature extremes T_{max} , and (2) Wetbulb temperature extremes TW_{max} . Up to now, there has been no investigation using EVT of the temperature extremes in southern Pakistan (Sindh). Thus, considering the need and relevance of the information such a study is necessary and timely. Therefore, we estimate the return levels of extreme daily maximum temperatures T_{max} and daily maximum wet-bulb temperatures TW_{max} over the different return periods during summer (May-September) in Sindh. We apply the POT method on the observational data of the nine weather stations provided by Pakistan Meteorological Department, and the ERA Interim re-analysis data of European Center for Medium range Weather Forecast (ECMWF) model for the corresponding grid points from 1980 to 2013. ERA Interim reanalysis data are generally very good at replicating trends in percentile-based measures of temperature extremes (Cornes and Jones, 2013). But it is in principle not obvious that ERA data can simulate well meteorological extremes, as reanalysis are constructed in such a way that typical conditions are well reproduced. This is why we look at how well ERA data performs in the target area against observations. If the ERA Interim dataset characterizes well the extremes, it could be an option for the regions within Sindh where no observational data is available. Furthermore, a standard bias correction is applied on the ERA Interim data to assess whether the return levels of extremes are better predicted after the rescaling. As described below, given the shortness of the datasets, it is appropriate at this stage to analyze the extremes without taking into considerations possible long-term trends.

The paper is organized as follows. In Section 2, the statistical modeling of extremes using the POT method is briefly described along with a description of the data used. Section 3 presents the main results of the POT analysis of the meteorological station observations, ERA Interim, and bias corrected ERA Interim daily

maximum temperature T_{max} and wet-bulb temperature TW_{max} data at nine locations, viz. Jacobabad, Mohenjodaro, Rohri, Padidan, Nawabshah, Hyderabad, Chhor, Karachi, and Badin. The performance of the ERA Interim and bias corrected ERA Interim in comparison to observations is also described in Section 3. All computations and graphics in this work are done using the R free open source statistical software, using the packages ismev and extRemes (see www.R-project.org and R Development core team 2015). Section 4 summarizes the major findings of the study and concludes our work.

2. Data and Methodology

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2.1 Meteorological station data

The daily maximum temperature and relative humidity data recorded at nine meteorological stations in Sindh from 1980 to 2013 are provided by the Pakistan Meteorological Department (see Table 2). We select nine stations, which contain a negligible amount of missing values after 1980, and are suitable for the POT analysis. An additional criterion is that only those stations are chosen where no changes occurred in measuring instruments during the last 33 years (Brunetti et al., 2006). None of the station data shows gaps with a duration longer than two days, which are treated by replacing the missing value with the average of the two previous values.

The temperature data are discretized unevenly with intervals up to 1 degree Celsius. Deidda and Puliga (2006) proposed a Monte Carlo approach for addressing this issue. They showed that finite resolution in precipitation data affects the convergence of parameter estimation in the extreme value analysis. They suggested generating many synthetic datasets by adding numerical noise to the original data, and then providing the best estimate of the parameters of the extreme value distributions by averaging over all the best fits obtained in each synthetic dataset. Following their suggestion, we produce high resolution data to compensate the effect of discretization and thus to improve the convergence of the estimator. In order to convert the temperature readings to higher resolution, we add a uniform random variable in the interval [-0.5, 0.5]. The main property of this noise is that round (T+r) = T, where T is the temperature with 1-degree resolution and 'round' is the numerical function, which maps the interval [T-0.5, T+0.5] to T. Thus, adding the noise does not perturb the information content of the observations. This procedure is applied to all temperature data, irrespective of the actual resolution, and replicated 100 times using a Monte Carlo approach. For each synthetic dataset, we perform the statistical best fit described later in the paper and then average the results. We check the influence of this noise parameterization and find no significant bias in the return level estimates. The advantage of adding a noise is to avoid the spurious statistical effects associated to the presence discrete values assigned to the temperature readings. Using the described bootstrap method we reduce such problem without biasing the data.

2.2 ERA Interim re-analysis data

The gridded daily maximum temperature and relative humidity data of ERA Interim re-analysis is obtained from the ECMWF Public Datasets web interface (http://apps.ecmwf.int/datasets/). The ERA Interim is generated by the European Center for Medium range Weather Forecast (ECMWF) model with resolution $0.75^{\circ} \times 0.75^{\circ}$ (Dee et

al., 2011). The gridded data are then extracted at the closest grid points of all stations, for the period 1980-2013. The latitude and longitude of the ERA Interim stations are displayed in Table 2.

The extreme temperatures analysis is restricted to the summer season (May-September) over a period of 33 years. We have tested that trends are not significant in such a short time interval. One of the main requirements for performing the POT analysis is assuming the stationarity of the time series. Therefore, as in Bramati et al. (2014), the Augmented Dickey Fuller (ADF) test of stationarity is performed on all time series (Dickey and Fuller, 1979). In all cases we find no sign of long-term correlations in the data. Short-term correlations (daily time scale) typically lead to clusters of extreme values and are studied by computing the extremal index θ in all time series and treated using the associated standard declustering technique (see more details in Section 2.4).

2.3 Wet-bulb temperature calculations

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The wet-bulb temperature measures the heat stress better than other existing heat indices, because it establishes the clear thermodynamic limit on heat transfer that cannot be overcome by adaptations like clothing, activity and acclimatization (Pal and Eltahir 2015, Sherwood and Huber, 2010). Here, we use an empirical equation developed by Stull (2011) to measure the wet-bulb temperature.

$$TW = T \arctan (\alpha_1 \sqrt{RH + \alpha_2}) + \arctan(T + RH) - \arctan(RH + \alpha_3) + \alpha_4 (RH)^{\frac{3}{2}} \arctan(\alpha_5 RH) - \alpha_6$$
20 (1)

where TW is the wet-bulb temperature [°C], T is the temperature [°C], and RH is the relative humidity [%]. This relationship is based on an empirical fit, as in Stull (2011), where the coefficient values are $\alpha_1 = 0.151977$, $\alpha_2 = 8.313659$, $\alpha_3 = -1.676331$, $\alpha_4 = 0.00391838$, $\alpha_5 = 0.023101$, and $\alpha_6 = 4.686035$. Equation (1) covers a wide range of relative humidity and air temperatures with an accuracy of 0.3°C.

2.4 Peaks over Threshold

In order to determine the return levels of extreme maximum temperatures and maximum wet-bulb temperatures, the peaks over threshold approach is applied to the data obtained from the meteorological stations in Sindh, and from the ERA Interim archive.

Multi-occurrence is an important characteristic of extreme climatic events and is referred to as clustering. Clusters are consecutive occurrences of above threshold events. It is important to post process the clustered extremes in order to take into account the assumption of weak short time correlation between extreme events, which is crucial for our statistical analysis. We have treated the clusters using the concept of Extremal Index (EI) (see Newell, 1964, Loynes, 1965, O'Brien, 1974, Leadbetter, 1983, Smith, 1989, Davison and Smith, 1990). The Extremal Index θ measures the degree of clustering of extremes. It ranges between 0 and 1, (θ = 0 means strong clustering and dependence, θ = 1 absence of clusters and independence). Leadbetter (1983) interprets $1/\theta$ as the mean number of exceedances in a cluster.

The extremal index θ can be estimated in two separate ways. Here, we apply the 'intervals estimator' automatic declustering by Ferro and Segers (2003). A positive aspect of this method is that it avoids the subjective choice of cluster parameters. The main ingredient is the use of an asymptotic result for the times between threshold exceedances. The exceedance times are split into two types, a set of vanishing intra-exceedance times within the clusters, and an exponentially distributed set of inter-exceedance times between clusters. The method is iterative, starting with largest return times and stops when a limit for the inter-exceedance times is reached. The standard errors of the estimated parameters is obtained by a bootstrap procedure. In this study, once we select appropriate value for the threshold (see below) the extremal index value is ≤ 0.5 in all the considered time series. Therefore, it is necessary to decluster the extremes by choosing the largest event in each cluster, before fitting it to the GPD.

As mentioned before, we use as statistical model for the exceedances over threshold the Generalized Pareto Distribution (GPD), which is characterized by two parameters, the shape ξ and the scale σ . The GPD for exceedances x - u of a random variable x reads as

$$G(x) = 1 - \left[1 + \xi \left(\frac{x - u}{\sigma}\right)\right]^{-\frac{1}{\xi}}$$
 $(x > u, \xi \neq 0)$, (2)

where u is the threshold. The shape parameter ξ determines the tail behavior while the scale parameter σ measures the variability. For a negative shape parameter, $\xi < 0$, the distribution is bounded (Weibull distribution), for vanishing shape parameter, $\xi = 0$, the distribution is exponential, and for a positive shape parameter, $\xi > 0$, the distribution has no upper bound (Pareto distribution).

In particular, for a negative shape parameters ξ <0 the GPD has the upper bound

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$$A_{max} = u - \sigma/\xi$$

$$G(x) = 0 \qquad (x > A_{max}, \xi < 0)$$

where A_{max} is an absolute maximum (Lucarini et al., 2014). In general, the best estimate for the two parameters shape ξ and scale σ depend on the threshold u (Coles, 2001). The choice of the optimal threshold for performing statistical inference from a time series is crucial. Choosing a very large value for u reduces the number of exceedances to a few values, inflating the variance of the estimators, so that the analysis is unlikely to yield any useful results. On the other hand, choosing a too small value for u would violate the asymptotic nature of the model, with a possible biased estimation and wrong model selection (Coles, 2001), see details later in Section 3.1. The shape ξ , the scale σ and the return levels are estimated using the Maximum Likelihood Estimator (MLE) using the R software (R Development core team 2015), which also provides an estimate of the standard error of the estimates.

Additionally, we wish to investigate the N - years return levels x_N , which are exceeded on the time scale of N years (Coles, 2001) and can be expressed as

$$x_N = u + \frac{\sigma}{\xi} \left[(N n_y \zeta_u)^{\xi} - 1 \right], \tag{4}$$

where N represents the return period in years, n_y is the number of observations per year, ζ_u is the probability of an individual observation exceeding the threshold u, the shape parameter is ξ and the scale parameter is σ .

5 2.5. Bias Correction Method

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A simple bias correction is applied to each ERA Interim time series through a rescaling that adjust the first two moments (mean and variance) to the sample moments calculated on the corresponding observations. Therefore, the bias correction is applied to the entire time series and it is not tailored to the extreme events only. The idea is to check whether by adjusting the properties of the bulk of the statistics we improve considerably the skill of the ERA Interim dataset in describing extreme events. The bias corrected ERA Interim time series *x* is expressed as

$$x = \bar{z} + \frac{y_{ERA} - \bar{y}}{\sigma_y} \sigma_z \tag{5}$$

where y_{ERA} is the ERA Interim time series, \bar{y} and σ_y its mean and standard deviation, whereas \bar{z} and σ_z are the mean and standard deviation of the meteorological station temperatures.

3. Results and Discussion

3.1 Threshold Selection

The threshold selection is the first step in a POT analysis. One needs to test whether the asymptotic regime is reached, i.e. whether we are choosing true extremes. This can be investigated by checking whether the best fits of the shape parameter ξ and the modified scale parameter σ*= σu – ξu are stable with respect to increases in the chosen value of u (Sacrrott and MacDonald, 2012). The optimal threshold u is selected as the lowest value where the two parameters are invariant in order to reach the asymptotic limit (Coles, 2001 and Furrer et al., 2010). This choice allows for having as many data as possible for performing the statistical inference, thus having lower variance for the estimators of the parameters. Figure 2 shows the parameter stability plots of the T_{max} reading for Karachi, as an example to explain the threshold selection procedure.

In addition to diagnostic plots of the modified scale parameter σ^* and the shape parameter ξ , the mean residual life plot is used to select the appropriate threshold for the POT analysis (Davison and Smith, 1990). The idea is to select the lowest value of the threshold when the plot is approximately linear. In the case of the Karachi,data for T_{max} , the plot appears to be linear and stable for $u = 36^{\circ}\text{C}$, indicating u = 36 as the most suitable threshold for the POT analysis (Figure 3). We observe that the 90% quantile is an appropriate threshold for all the station data, as well as the ERA interim datasets, and for both T_{max} , and TW_{max} .

35 3.2 GPD Fit

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The goodness of fit is evaluated by Quantile-Quantile (Q-Q) plots and hypothesis testing. The Q-Q plot analysis is performed for the stations observed, the ERA Interim, the bias corrected ERA Interim daily T_{max} and TW_{max} .

The Q-Q plots of the observed T_{max} show that the GPD fits well in most stations. However, in a few stations like Jacobabad, Mohenjo-daro, Padidan and Chhor the empirical values show slight deviation from the modeled values. In spite of minor deviations at some stations, still most of the exceedances are well fitted by the model. The Q-Q plots of the observed TW_{max} also fits well to the model in all stations.

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The Q-Q plots of the ERA Interim T_{max} and TW_{max} reveals substantial differences with respect to the GPD. The empirical values of the higher quantiles are deviating from the theoretical quantiles in all stations. However, if the higher quantiles are disregarded, then stations like Jacobabad, Mohenjo-daro, Rohri, Padidan, Nawabshah, Chhor, and Badin fits very well with the model. The Q-Q plots of the bias corrected ERA Interim T_{max} , and TW_{max} show better results than the ERA Interim. We notice that the T_{max} of the ERA Interim and bias corrected ERA Interim fits better than the TW_{max} if the higher quantiles are ignored, indicating the bias procedure is, as expected, unable to treat correctly the statistics of the largest events.

In order to assess the goodness-of-fit, we apply the Kolmogorov-Smirnov (K-S) test and Anderson-Darling (A-D) test to the data of meteorological stations, ERA Interim, bias corrected ERA Interim T_{max} and TW_{max} . The p-values indicate a good performance of the fit procedure. Table 3 shows the results of the K-S and A-D statistics of the T_{max} and TW_{max} in all the data sets.

3.3 Parameter Estimates

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Here, we analyze the shape parameter ξ , the scale parameter σ , and threshold u for all considered datasets. The standard errors of the shape ξ and the scale σ parameters are given in Table 4. The spatial distribution of the shape parameter ξ and the scale parameter σ of the GPD in Sindh are shown in Figure 4. The shape parameters ξ are negative in all datasets at all stations. This is hardly surprising, as meteorological and physical processes make sure that the temperature cannot grow locally without control. Figure 4 displays the bias corrected ERA Interim results only. The observed T_{max} shape parameters ξ are between -0.418 to -0.223, and for TW_{max} within -0.323 to -0.177. The bias corrected ERA Interim T_{max} shape parameters ξ range from -0.305 to -0.002, and TW_{max} are between -0.18 to -0.01. The agreement in the values of the shape parameters in the observations and simulations means that the ERA dataset captures an important aspect of extremal behavior. This is in principle a non-trivial result, as reanalysis are constructed in such a way that typical conditions are well reproduced.

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The scale parameters σ of the observed T_{max} range from 2.08 to 2.76, and the TW_{max} are in 1.86 to 2.76. In the ERA Interim analysis, the scale parameter σ of T_{max} is between 1.00 - 1.95, and TW_{max} in 0.74 -1.75. We observe a difference in the scale parameters of both the observed, ERA Interim T_{max} and TW_{max} . We find that, unsurprisingly, the scale parameters of the bias corrected ERA Interim data are much closer to those estimated for T_{max} and TW_{max} using the station data. In the bias corrected ERA Interim T_{max} the scale parameters σ are in 1.50 - 2.75, while for TW_{max} are in a range 1.40 - 2.40 (Figure 4). All the temperature scale parameters are in degree Celsius.

3.4 Absolute Maxima

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Once the shape parameters ξ , the scale parameters σ , and the thresholds u are determined, it is possible to compute the theoretical absolute maxima using Eq. (3) (Section 2.4). Theoretical absolute maxima can be compared with the observed ones for each station to better understand whether our fits are in agreement with the observed data. The daily maximum temperature T_{max} and the maximum wet-bulb temperature TW_{max} (station data, the ERA Interim, and the bias corrected ERA Interim) have negative shape parameters ξ at all stations. This means that according to Eq. (2) in section 2.4, the probability distribution function (pdf) is bounded by the maximum values. These maximum values are the theoretical upper limits predicted by the GPD fit. The analysis shows that the observed absolute maxima T_{max} and TW_{max} at all stations of the three data sets are below the theoretical absolute maximum, as expected (Figure 5). This gives us confidence on the quality of our fit. The following piece of information can also be derived: assume that one observes in the future an extreme event larger than the maximum inferred in the present dataset; this may suggest some non-stationarity in the most recent portion of the dataset.

3.5 Return Levels

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The return levels (RLs) are computed considering various return periods (2, 5, 10, 20, 50, 100-year). The return level plots of the stations observed, the ERA Interim, the bias corrected ERA Interim daily maximum temperature T_{max} and daily maximum wet–bulb temperature TW_{max} are displayed in Figures 6 and 7. The values of the RLs follow the north-south gradient of the climatic mean temperatures. The northern part of the Sindh (Jacobabad, Mohenjo-daro, Rohri, Padidan, and Nawabshah) are hotter than the southern part (Hyderabad, Chhor, Karachi, and Badin).

The 2, 5, 10, 20, 50, 100-year RLs estimated in Sindh for station observed T_{max} at time reach over 50°C in Jacobabad, Mohenjo-daro, Padidan, Nawabshah, and over 45°C in Rohri, Hyderabad, Chhor, Karachi, Badin. The corresponding ERA Interim T_{max} return levels are at least 3°C to 5°C lower in all stations, while having correct representation of the geographical variability of the field. As example, the RLs of 42°C at Badin has a 3-year return period in the observations T_{max} , but a 30-year return period in ERA Interim (Figure 6).

The RLs of TW_{max} are above 35°C in all meteorological stations. As for the ERA Interim, the RLs of TW_{max} are greater than 30°C for all the stations except Karachi, which has RLs less than 30°C. Here, we see again that the RLs of the ERA Interim TW_{max} are lower than the RLs of station TW_{max} . Going again to the Badin stations, the 4-year return period observed for TW_{max} is 38°C, while the ERA Interim dataset show the same RL in a 15-year return period (Figure 7).

The bias corrected ERA Interim T_{max} and TW_{max} , show some improvements in the RLs at all stations. When looking at the Nawabshah, Hyderabad, Karachi, and Badin stations, the RLs agree with those obtained from the station data in the range 5-100 years, while disagreements exist in the range 2-5 years. In the rest of the stations, the bias corrected data RLs are closer to those of the station data, yet not statistically compatible with them. When looking at the wet-bulb temperature TW_{max} analysis, the RLs of the bias corrected ERA Interim show some overlap with the those derived from station observations in Mohenjo-daro, Hyderabad, Chhor, and while no overlap is found in the other stations. One understands that the proposed simple bias correction methods

improves the quality of the representation of extremes by ERA Interim, but many discrepancies remain (Figures 6 and 7).

We also plot the station and bias corrected ERA Interim T_{max} , and TW_{max} return levels spatially for the 5, 10, 25 and 50-year return periods (Figures 8 and 9), as a detailed spatial overview of the temperature extremes in Sindh might be of interest to the policy makers.

4. Summary and Conclusion

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The main objective of this study is the assessment of the return levels of the extreme daily maximum temperatures T_{max} and wet-bulb temperatures TW_{max} in southern Pakistan (Sindh). In addition, the performance of the ERA Interim TW_{max} is compared to the weather station TW_{max} to assess its ability to estimate temperature extremes in Sindh. Moreover, a simple bias correction is applied to the ERA Interim data to see whether correcting the first two moments of its statistics helps in improving its performance in representing temperature extremes.

The peaks over threshold method is applied to the daily T_{max} and TW_{max} data of nine stations and to the corresponding nearest ERA Interim temperature data. Standard declustering technique is applied to all time series to achieve the independence assumption of extremes. After testing the asymptotic statistical properties, the 90% quantile is found to be appropriate threshold choice for all the weather stations, the ERA Interim and the bias corrected ERA Interim maximum temperature and wet-bulb temperature. A Generalized Pareto Distribution (GPD) is fit to both T_{max} and TW_{max} for all three datasets. The results conclude that the shape parameter ξ is negative at all stations. The scale parameter σ estimated on weather station temperatures is much closer to the bias corrected ERA Interim estimates than the original ERA Interim data ones. The theoretical absolute maxima of the time series are higher than the observed absolute maxima in all stations. The Q-Q plots are used to assess the GPD fit, which results to be acceptable for both T_{max} and TW_{max} station data as compared to the ERA Interim data. However, the bias corrected ERA Interim data shows improved GPD fits than ERA Interim data.

The return levels (RLs) of T_{max} and TW_{max} are estimated for the 2, 5, 10, 25, 50, 100-year return periods in all datasets. The RLs of T_{max} estimated using the meteorological station temperatures are greater than 50°C in Jacobabad, Mohenjo-daro, Padidan, Nawabshah, and greater than 45°C in Rohri, Hyderabad, Chhor, Karachi and Badin. While the RLs of TW_{max} in station data are larger than 35°C in the entire Sindh, when using ERA Interim temperatures, they are estimated as greater than 45°C in Northern Sindh and greater than 40°C in southern Sindh. The differences in the RLs using the two datasets are between 3°C and 5°C for both shorter and longer return periods due to the minor variations in the shape and scale parameters. Although the ERA-Interim dataset does not capture well the magnitude of the extremes, but it provides a good representation of their spatial fields.

The bias corrected ERA Interim T_{max} and TW_{max} gives return levels closer to the meteorological stations observed ones than the original ERA Interim return levels at all stations. Although the bias corrected ERA Interim shows a good correspondence with the meteorological station data, statistically differences remain in most cases. Therefore, one must use more advanced bias correction method for analyzing extremes precisely.

The extremes of daily maximum wet-bulb temperature TW_{max} are estimated as above the human survivability threshold 35°C throughout the region, so the risk of hyperthermia is very high here. The most vulnerable people are those who are involve in the everyday outdoor activities like farming, fishing, building construction, athletes, elderly and infants can have heat strokes, dehydration etc. The human habitability in such a warm region is already at risk and one can expect that these issues will be worse in future climate conditions.

The crops are very sensitive to temperature variations, and even a rise of one degree Celsius can cause detrimental changes in the phenological stages of the crops (Hatfield and Preuger, 2015). Every crop has a certain limit to tolerate the temperature. When temperature exceeds this limit, the crop yield is drastically reduced. In summer, the temperature and humidity increase to an extent that there are high chances of a rapid pests spread in the crops. Sindh produces cotton, wheat, rice, mango, banana, and dates, so a correct estimate of temperature extremes is very important.

This clarifies that the biases between the station and the ERA Interim data are rather relevant when one wishes to address the impact of hot climatic extremes to human life and to active crop production in the region. It would be of primary importance to understand the physical reasons behind such inconsistencies, which makes it hard to use reasonably ERA without bias correction. Clearly, they might result either from a misrepresentation of local processes dominated by near surface processes (namely, heat and water fluxes), or from an inadequacy of the reanalysis in reproducing synoptic and sub-synoptic conditions responsible for extremely hot and humid conditions. This matter is surely worth investigating but is well beyond the scope of this paper.

This paper contains beneficial information regarding the assessment of the temperature extremes in Sindh, which could help the local administrations to prioritize the regions in terms of adaptations like preparation of baseline contingency plans for dealing with strong heat waves based on the current climatology. Such measures are not yet present in the territory and lead to many casualties each year. While the stationary analysis presented here has already relevance in terms of impacts for the public and private sector as it fills a research gap, and is statistically motivated by the short duration of the observational dataset (33 years). Indeed, it seems relevant to investigate time dependency in the temperature extremes. We might consider using the centennial NCEP reanalysis (Compo et al., 2011) and using suitable bias correction procedures. We propose to repeat this analysis in GCMs (CMIP5, CMIP6) and RCMs (CORDEX) to study the properties of extremes. All models use re-analysis as input, and generate information of extremes, which involves biases that if not corrected, can lead to significant errors in prediction of present and future extremes. Therefore, in order to reduce the uncertainties in impact assessment, it is necessary to improve the re-analysis before using it in GCMs and RCMs.

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We would like to thank Climate KIC, for funding this research. This publication is a part of a Climate KIC project "Extreme Events in Pakistan: Physical processes and impacts of changing climate", which belongs to the adaptation services platform of the Climate KIC. Thanks to Pakistan Meteorological Department (PMD) and the

European Center for Medium range Weather Forecast (ECMWF) for providing the datasets. The R development core team (2015) is acknowledged for providing statistics packages. We would like to thank the DFG Cluster of Excellence CliSAP for partially supporting this research activity.

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Table 1. Monthly mean climatic characteristics of all nine stations from 1980-2010.

Stations						Mean T	empera	ture (°C)				
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	Annual
Jacobabad	15.2	18.2	24	30.5	35.6	37	34.8	33	31.4	27.8	22.3	16.7	27
Mohenjo-	13.9	16.7	23	29.1	34.1	35	33.9	32.9	30.9	26.7	21.1	15.9	25.9
daro													
Rohri	15.6	18.2	23.6	29.8	34.5	35.6	33.9	32.3	31.2	27.6	22.1	16.9	26.4
Padidan	14.8	17.7	23.5	29.9	34.4	35.5	33.7	32.1	31	27.5	22.4	16.4	26.5
Nawabshah	15.4	18	24	29.8	34.5	35.6	34	32.3	31.5	28	22.4	16.9	26.7
Hyderabad	18	21	26.2	30.9	33.3	34	32.4	31.1	31	29.6	24.8	19.6	27.6
Chhor	16.5	19.5	25	30.1	33.5	33.7	31.6	30.1	30.1	28.2	22.6	17.9	26.3
Karachi	18.6	21.2	25.4	28.9	31.1	31.9	30.5	29.2	29.5	28.9	24.6	20.4	26.4
Badin	17.5	20.5	25.8	30.1	32.6	32.8	31	29.6	29.6	28.7	24	19	26.6
Stations					N	Ainimun	1 Tempe	rature (°	PC)				
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	Annual
Jacobabad	7.9	10.9	16.6	22.4	27.4	29.8	29.3	28.4	26.3	20.5	14.3	8.9	19.9
Mohenjo-	4.7	7.9	13.3	18.9	24	27.4	27.9	27	24.7	18.2	11.8	7.3	17.3
daro													
Rohri	8.3	10.8	15.9	21.7	26.1	27.7	27.1	26	24.4	19.9	14.2	9.6	18.7
Padidan	6.5	8.9	14.5	20.2	24.7	27	26.9	25.8	23.7	18.3	12.4	7.6	17.8
Nawabshah	6.3	8.7	14.2	19.4	24.6	27.3	27.2	25.9	23.8	18.4	12.4	7.8	17.9
Hyderabad	11.4	13.9	18.8	22.8	26.1	27.9	27.6	26.5	25.4	22.5	17.4	13	21.1
Chhor	5.9	8.9	14.8	20.3	24.8	26.9	26.5	25.3	23.9	18.7	11.8	7	17.6
Karachi	11.5	14	18.6	23	26.6	28.3	27.6	26.3	25.6	21.9	16.8	12.7	20.7
Badin	9.9	12.6	17.9	22.3	25.7	27.6	27.1	26	25	22.1	16.5	11.4	20.2
Stations		Maximum Temperature (°C)											
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	Annual
Jacobabad	22.6	25.6	31.4	38.6	43.9	44.4	40.2	37.6	36.8	35.1	30.3	24.4	34.1
Mohenjo-	23.1	26.2	32.1	38.7	43.8	44.2	40.9	38.7	37.5	35.2	30.5	24.8	34.5
daro													
Rohri	22.6	25.6	31.2	38.1	43	43.5	40.5	38.3	37.8	35.2	30	24.3	34
Padidan	23.1	26.4	32.2	39.4	43.9	44.1	40.6	38.4	38.3	36.3	31.1	25.3	34.8
Nawabshah	24.5	27.9	33.8	40.2	44.2	43.9	40.7	38.8	39	37.7	32.3	26.1	35.5
Hyderabad	24.7	28.1	33.7	38.8	41.3	40	37.2	35.6	36.3	36.7	31.9	26.2	34.1
Chhor	26.9	29.9	35.2	40	42	40.6	36.8	34.9	36.3	37.6	33.5	28.7	35
Karachi	26.3	28.4	32.2	34.7	35.5	35.4	33.3	32.1	33.2	35.5	32.5	28.2	32
Badin	25.2	28.3	33.7	37.8	39.4	37.9	34.9	33.2	34.2	35.2	31.4	26.5	32.9

Table 2. Code, Name, Geographic coordinates and Altitude of the stations.

Code	Name	PI	MD weather statio	ERA-Interim stations		
	Name	Latitude	Longitude	Altitude (m)	Latitude	Longitude
JCB MJD	Jacobabad Mohenjo-daro	28° 18'N 27° 22'N	68° 28'E 68° 06'E	55 52.1	28°4'N 27°5'N	68°15'E 67°75'E
RHI	Rohri	27° 40'N	68° 54'E	66	27°75'N	69°25'E
PDN	Padidan	26° 51'N	68° 08'E	46	26°8'N	68°5'E
NWB	Nawabshah	26° 15'N	68° 22'E	37	26°25'N	68°0'E
HYD	Hyderabad	25° 23'N	68° 25'E	40	25°5'N	68°15′E
CHR	Chhor	29° 31'N	69° 47' E	5	25°3'N	69°6'E
KHI	Karachi	24° 54'N	67°08' E	21	25°2'N	67°5′E
BDN	Badin	24° 38'N	68° 54'E	10	24 °75'N	68°65'E

Table 3. Results of the Kolmogorov-Smirnov Goodness of fit test and Anderson-Darling test between empirical and GPD fits.

			Observe	d Tmax							
Test	Null	P-value									
Statistics	Hypothesis	JAC	MJD	RHI	PDN	NWS	HYD	CHR	KHI	BDN	
Kolmogorov Smirnov	Equality of probability distribution	0.947	0.340	0.996	0.139	0.941	0.385	0.928	0.306	0.666	
Anderson Darling	Equality of probability distribution	0.553	0.978	0.654	0.857	0.157	0.649	0.233	0.869	0.145	
			ERA Inter	rim T <i>max</i>	;						
Test	Null		<i>P</i> -value								
Statistics	Hypothesis	JAC	MJD	RHI	PDN	NWS	HYD	CHR	KHI	BDN	
Kolmogorov Smirnov	Equality of probability distribution	0.169	0.125	0.553	0.456	0.322	0.187	0.419	0.456	0.332	
Anderson Darling	Equality of probability distribution	0.355	0.263	0.165	0.587	0.615	0.398	0.266	0.687	0.425	
		Bias co	orrected EF	RA Interii	m T <i>max</i>						
Test	Null			j	<i>P-</i> value			•			
Statistics	Hypothesis	JAC	MJD	RHI	PDN	NWS	HYD	CHR	KHI	BDN	
Kolmogorov Smirnov	Equality of probability distribution	0.452	0.4729	0.197	0.489	0.269	0.137	0.158	0.243	0.312	
Anderson Darling	Equality of probability distribution	0.352	0.315	0.235	0.270	0.335	0.289	0.216	0.390	0227	
			Observed	TWmax							
Test	Null	P-value									
Statistics	Hypothesis	JAC	MJD	RHI	PDN	NWS	HYD	CHR	KHI	BDN	
Kolmogorov Smirnov	Equality of probability distribution	0.981	0.111	0.341	0.226	0.457	0.545	0.441	0.385	0.211	
Anderson Darling	Equality of probability distribution	0.623	0.745	0.587	0.884	0.199	0.123	0.789	0.669	0.473	
			ERA Interi	m TW <i>ma</i>	x						
Test	Null				<i>P</i> -value						
Statistics	Hypothesis	JAC	MJD I	RHI	PDN N	NWS	HYD	CHR	KHI	BDN	
Kolmogorov Smirnov	Equality of probability distribution	0.712	0.564	0.955	0.425	0.258	0.134	0.856	0.497	0.222	
Anderson Darling	Equality of probability distribution	0.236	0.474	0.516	0.219	0.356	0.117	0.537	0.464	0.613	
		Bias co	rrected ER	A Interin	ı TW <i>max</i>						
Test	Null	IAG	MID	-	P-value	NINKIC	HIVD.	CHE	1/111	DDE	
Statistics	S Hypothesis Equality of probability	JAC	MJD	RHI	PDN	NWS	HYD	CHR	KHI	BDN	
Kolmogorov		0.268	0.688	0.127	0.372	0.268	0.229	0.591	0.582	0.478	
Smirnov Anderson	distribution Equality of probability	0.200	0.000								

Table 4. Estimated parameters shape ξ , scale σ and standard error $\Delta\xi$ of all the data sets.

Station observed T _{max}												
Estimates	JCB	MJD	RHI	PDN	NWB	HYD	CHR	KHI	BDN			
Shape ξ	-0.3875	-0.2550	-0.4182	-0.3261	-0.3323	-0.3292	-0.3108	-0.2225	-0.3292			
Standard Error $\Delta \xi$	0.0317	0.0226	0.0226	0.0218	0.0208	0.0312	0.0371	0.0341	0.0312			
Scale σ	2.7540	2.0819	2.3510	2.2144	2.1391	2.2286	2.5629	2.5685	2.2286			
Standard Error $\Delta \sigma$	0.1421	0.1040	0.1075	0.1076	0.1031	0.1166	0.1462	0.1444	0.1166			
ERA Interim T _{max}												
Estimates	JCB	MJD	RHI	PDN	NWB	HYD	CHR	KHI	BDN			
Shape ξ	-0.1959	-0.1788	-0.2076	-0.2185	-0.2135	-0.3380	-0.2850	-0.0376	-0.2514			
Standard Error $\Delta \xi$	0.0320	0.0348	0.0343	0.0287	0.0265	0.0316	0.0337	0.0508	0.0371			
Scale σ	1.4643	1.3230	1.3440	1.5045	1.5630	2.0656	1.8497	1.3303	2.0410			
Standard Error $\Delta \sigma$	0.0798	0.0739	0.0741	0.0788	0.0788	0.1082	0.0949	0.0908	0.1153			
			Bias Corr	ected ERA	Interim T	nax						
Estimates	JCB	MJD	RHI	PDN	NWB	HYD	CHR	KHI	BDN			
Shape ξ	-0.1959	-0.1788	-0.2076	-0.2185	-0.2135	-0.3380	-0.2850	-0.0376	-0.2514			
Standard Error $\Delta \xi$	0.0320	0.0348	0.0343	0.0287	0.0265	0.0316	0.0337	0.0508	0.0371			
Scale σ	1.9834	1.7918	1.8205	2.0382	2.1164	2.7980	2.3081	1.8016	2.7636			
Standard Error $\Delta \sigma$	0.1081	0.1001	0.1004	0.1068	0.1068	0.1467	0.1233	0.1229	0.1562			
			Stat	tion observ	ed TW _{max}							
Estimates	JCB	MJD	RHI	PDN	NWB	HYD	CHR	KHI	BDN			
Shape ξ	-0.1769	-0.1860	-0.2150	-0.2157	-0.2164	-0.3231	-0.2423	-0.2190	-0.1867			
Standard Error $\Delta \xi$	0.0383	0.0354	0.0347	0.0442	0.0266	0.0269	0.0347	0.0368	0.0322			
Scale σ	2.7590	2.0454	1.9600	2.0780	1.8572	2.3724	2.5126	2.3375	1.9032			
Standard Error Δσ	0.1596	0.1146	0.1084	0.1289	0.0938	0.1191	0.1380	0.1328	0.1055			
			E	RA Interim	TW_{max}							
Estimates	JCB	MJD	RHI	PDN	NWB	HYD	CHR	KHI	BDN			
Shape ξ	-0.0896	-0.0946	-0.0687	-0.1257	-0.1583	-0.1771	-0.0902	-0.0194	-0.1733			
Standard Error $\Delta \xi$	0.0379	0.0293	0.0327	0.0342	0.0313	0.0377	0.0357	0.0359	0.0378			
Scale σ	1.2879	1.2437	1.2311	1.4408	1.6104	1.6499	1.3423	0.6801	1.7886			
Standard Error Δσ	0.0748	0.0660	0.0676	0.0804	0.0875	0.0959	0.0760	0.0398	0.1028			
Bias Corrected ERA Interim TW _{max}												
Estimates	JCB	MJD	RHI	PDN	NWB	HYD	CHR	KHI	BDN			
Shape ξ	-0.08961	-0.0946	-0.06870	-0.12570	-0.15831	-0.17711	-0.09017	-0.01942	-0.17332			
Standard Error $\Delta \xi$	0.03786	0.02931	0.03275	0.03424	0.03134	0.03767	0.03571	0.03593	0.03782			
Scale σ	1.35674	1.64650	1.75852	1.49477	1.52013	2.05281	2.14609	1.39943	2.15299			
Standard Error $\Delta \sigma$	0.07878	0.08736	0.09651	0.08347	0.08254	0.11924	0.12145	0.08193	0.12370			

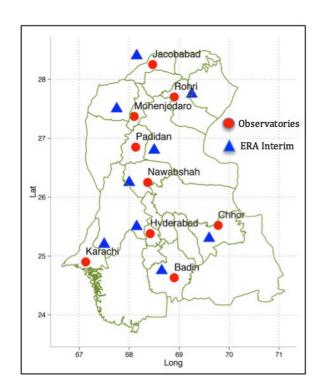


Figure 1: Study Domain (23.5 – 28.5° N , 66.5- 71.1°E)

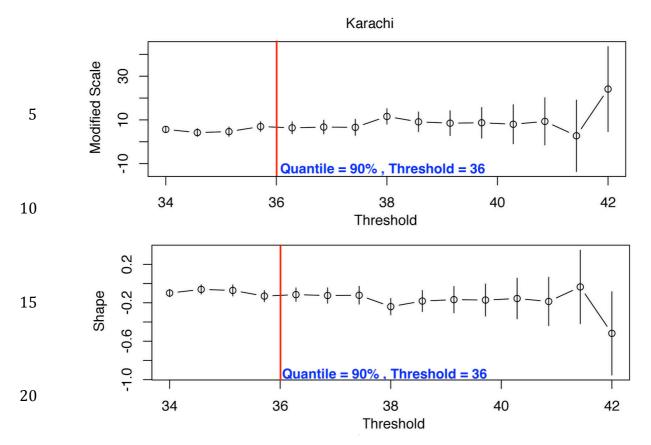


Figure 2. Modified scale (σ^*) and shape parameter (ξ) of the observed T_{max} (°C)Karachi. The red vertical lines represent the selected threshold according to the station quantiles.

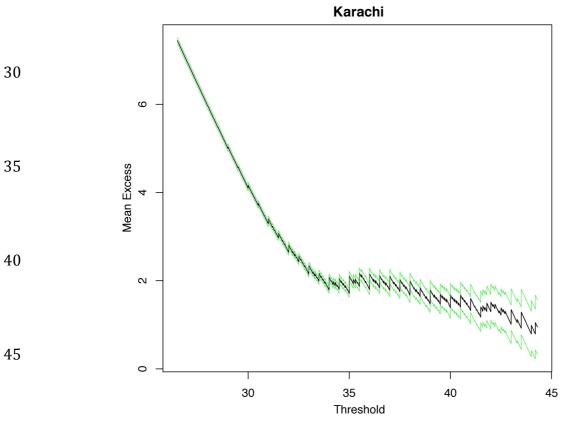


Figure 3. Mean residual life plot of the station observed $T_{\textit{max}}\,(^{\circ}\text{C})$ Karachi.

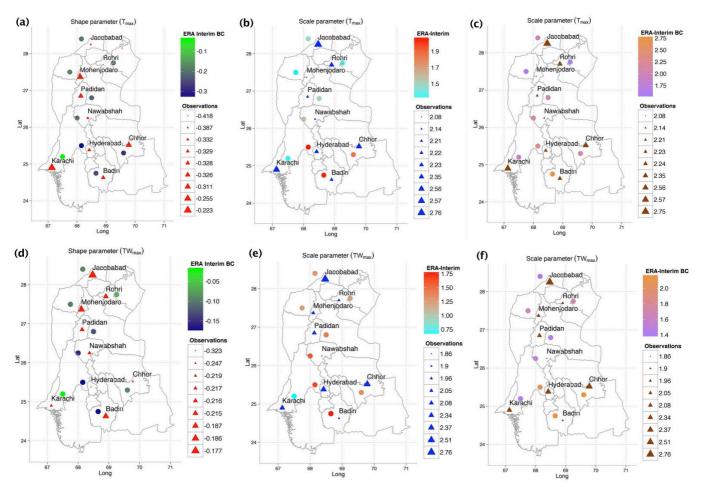


Figure 4. Spatial distribution of the shape parameters ξ and scale parameters σ of the station observed, ERA Interim, and bias corrected ERA Interim T_{max} (upper panel) and TW_{max} (lower panel) in degree Celsius.

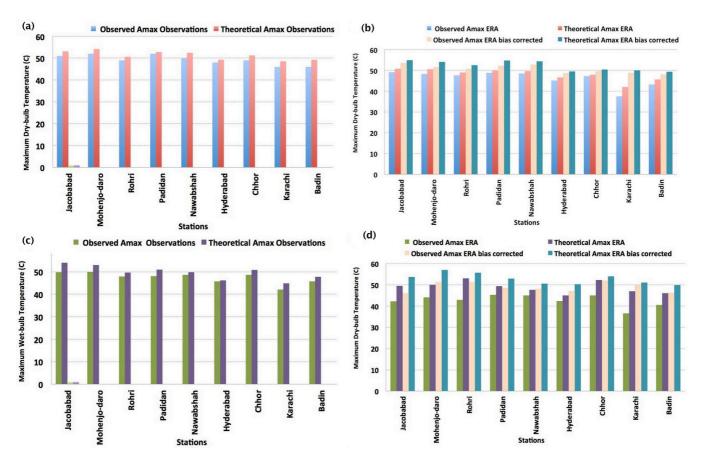


Figure 5. Absolute maxima A_{max} in degree Celsius (a) station observed T_{max} (b) ERA Interim and bias corrected ERA Interim T_{max} (c) station observed TW_{max} (d) ERA Interim and bias corrected ERA Interim TW_{max} .

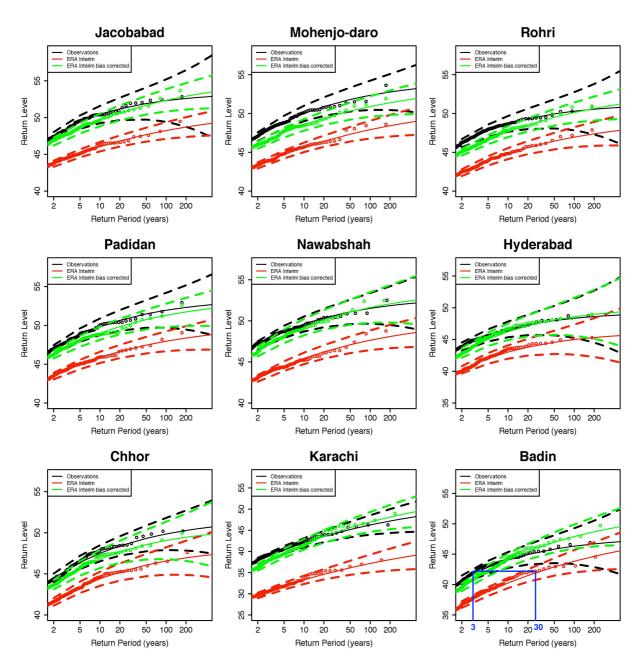


Figure 6. Return level plots of the station observed T_{max} (black), ERA Interim T_{max} (red), and bias corrected ERA Interim T_{max} (green) in degree Celsius. The blue line is to show a difference in the observed and ERA Interim RLs.

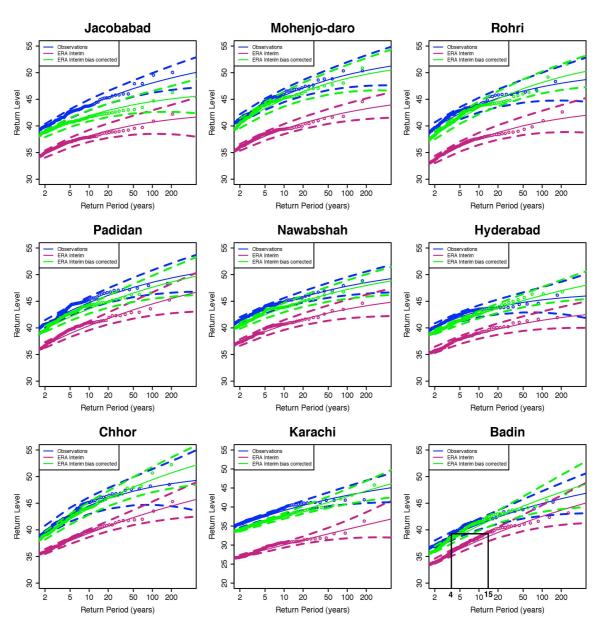


Figure 7. Return level plots of the station observed TW_{max} (blue), ERA Interim T_{max} (pink), and bias corrected ERA Interim T_{max} (green) in degree Celsius. The black line is to show a difference in the observed and ERA Interim RLs.

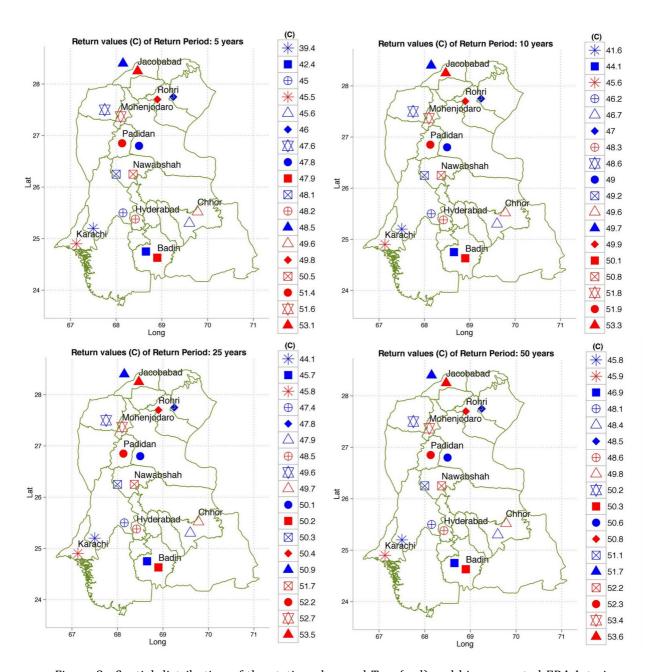


Figure 8. Spatial distribution of the station observed T_{max} (red) and bias corrected ERA Interim T_{max} (blue) return levels in degree Celsius corresponding to return periods of 5, 10, 25 and 50 years in southern Pakistan.

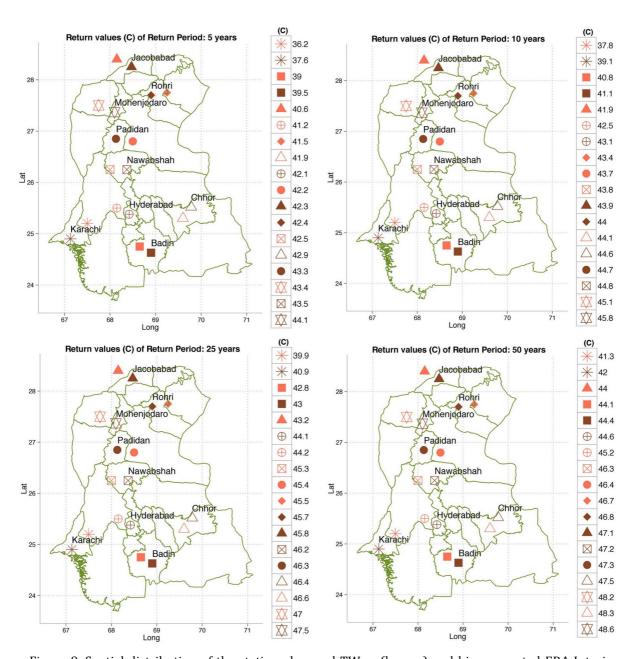


Figure 9. Spatial distribution of the station observed TW_{max} (brown) and bias corrected ERA Interim TW_{max} (orange) return levels in degree Celsius corresponding to return periods of 5, 10, 25 and 50 years in southern Pakistan.